

Sub-national violent conflict forecasts for sub-Saharan Africa, 2015-2065, using climate-sensitive models

Frank D W Witmer, Department of Computer Science & Engineering, University of Alaska Anchorage

Andrew M Linke, Department of Geography, University of Utah

John O'Loughlin, Institute of Behavioral Science and Department of Geography, University of Colorado Boulder

Andrew Gettelman, Climate and Global Dynamics Division, National Center for Atmospheric Research

Arlene Laing, Cooperative Institute for Research in the Atmosphere, Colorado State University

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Abstract

How will local violent conflict patterns in sub-Saharan Africa evolve until the middle of the 21st century? Africa is recognized as a particularly vulnerable continent to environmental and climate change since a large portion of its population is poor and reliant on rain-fed agriculture. We use a climate-sensitive approach to model sub-Saharan African violence in the past (geolocated to the nearest settlements) and then forecast future violence using socio-political factors such as population size and political rights (governance), coupled with temperature anomalies. Our baseline model is calibrated using 1° gridded monthly data from 1980-2012 at a finer spatio-temporal resolution than existing conflict forecasts. We present multiple forecasts of violence under alternative climate change scenarios (optimistic and current global trajectories), of political rights scenarios (improvement and decline), and population projections (low and high fertility). We evaluate alternate shared socioeconomic pathways (SSPs) by plotting violence forecasts over time and by detailed mapping of recent and future levels of violence by decade. The forecasts indicate that a growing population and rising temperatures will lead to higher levels of violence in sub-Saharan Africa if political rights do not improve. If political rights continue to improve at the same rate as observed over the last three decades, there is reason for optimism that overall levels of violence will hold steady or even decline in Africa, in spite of projected population increases and rising temperatures.

Keywords: disaggregated violence, multilevel models, governance, socioeconomic pathways, population projections, environmental change

Corresponding author: fwitmer@alaska.edu

Introduction

How will the local patterns of violent conflict in sub-Saharan Africa change through the middle of the 21st century? Africa is recognized as a particularly vulnerable continent to environmental and climate change since a large portion of its population is poor and is reliant on rain-fed agriculture. Our research uses a climate-sensitive approach to model sub-Saharan African violence in the past and then forecasts violence trends using key explanatory variables that have been shown to influence conflict. Forecasting the future is an exercise fraught with uncertainty but valuable when used as a tool to explore the outcomes of different, plausible, future political, social and climate scenarios. There is a growing literature on scenario development in the context of environmental decision making (Mahmoud et al., 2009; Steinitz, 2003) and climate change (O'Neill et al., 2014), and we adopt this general approach to explore shared socioeconomic pathways (SSPs) for violent conflict in sub-Saharan Africa. In particular, we consider future scenarios that vary fertility assumptions, governance trends and climate projections.

Conflict forecasting in the study of environmental change and conflict

Predictions are increasingly used in conflict research and are based on two general motivations. One approach uses predictions (either in- or out-of-sample validation) to assess the influences of independent variables upon an outcome of interest (Ward, Greenhill & Bakke, 2010; O'Loughlin, Linke & Witmer, 2014; O'Loughlin et al., 2012; Wischnath & Buhaug, 2014). Another uses modeling and simulation techniques to forecast observed trends into the future. Interest in this second kind of analysis has been growing in recent years (Hegre, et al., 2013; Hegre, et al. 2016; Hegre, et al. 2017; Ward et al., 2013; Blair, et al. 2017; Beger, Dorff & Ward, 2014; O'Brien, 2010; Schneider, Gleditsch & Carey, 2011; Schrod, Yonamine & Bogazzi, 2013). We are primarily motivated by this second type of forecasting, though initial steps in our research rely on model validation using observed data for recent years.

We forecast conflict using multiple future climate scenarios to address a question of great interest to academic and policy communities: how does climate variability and environmental stress lead to violent conflict? In doing so, we are contributing to a growing body of literature published in *Science* (Hsiang, Burke & Miguel, 2013), *Nature* (Hsiang, Meng & Cane, 2011), *Proceedings of the National Academy of Sciences* (Buhaug, 2010; Burke et al. 2009; O'Loughlin, Linke & Witmer, 2014; O'Loughlin et al., 2012; Schleussner et al. 2016) and *Global Environmental Change* (Böhmelt, 2014; Ide et al., 2014; Linke et al., 2015). Similarly, journals in the fields of political geography (De Juan, 2015), development

economics (Maystadt et al., 2015), climate studies (Theisen, Gleditsch & Buhaug, 2013), and peace research (Koubi et al., 2012) contain valuable contributions that consider particular social conditions that enable climatological extreme events to increase societal conflict. Robust debate continues about the relative importance of climate factors on conflict with Buhaug et al. (2014) arguing that Hsiang, Burke & Miguel (2013) overreach in their conclusion that deviations from normal temperatures increase the risk of conflict across multiple temporal and spatial scales.

Our contribution to the climate-conflict body of research therefore moves beyond the study of *only* the past and present. To that end, we consider two questions in the context of the climate-conflict literature. First, what will future patterns of conflict look like? Second, what factors are driving those future patterns? None of the research cited above includes substantial engagement with projected climate scenarios though generalized statements have been made. Hsiang, Burke & Miguel (2013: 1235367-1) argue, for example, “because locations throughout the inhabited world are expected to warm two to four standard deviations by 2050, amplified rates of human conflict could represent a large and critical impact of anthropogenic climate change.” But *where* would we observe such a spike in violence and under what circumstances? Answering these questions is a logical extension of existing research and is a key goal in this study. We quantify expectations about violent conflict patterns in geographically disaggregated predictions for sub-Saharan Africa until 2065.

In modeling and forecasting violence across all of sub-Saharan Africa at a 1° (degree) grid resolution, we explicitly consider the effects of temperature variability, while controlling for temporal reporting bias in the coding of our violent event data. Additionally, we account for a number of key social and political variables that have known associations with violence. The incorporation of these effects into regression analyses has been debated; while some encourage a thorough and dedicated effort to capture the specific geographical contexts and social settings of violence (e.g. Raleigh, Linke & O’Loughlin, 2014), others do not explicitly include these possible influences in statistical estimations even though data are available (e.g. Hsiang, Miguel & Burke, 2013). By including these socio-political factors with temperature variability, we are able to evaluate the relative contribution of each.

In this paper, we consider especially the role of political settings within which conflicts emerge. There is strong evidence that poor governance within countries – for example, institutions that are discriminatory or exclusionist against certain groups – contributes to the risk of coup attempts, secessionist movements, and scattered terrorist attacks. With this understanding, testing the relationship between rising temperatures and violence should be only one piece of a larger puzzle; identifying if, when and where temperature fluctuations might lead to conflict in different settings is the

key aim. Instead of making a single forecast of conflict into the future based on specific temperature estimates, we extend such an exploration by adding alternative scenarios for social indicators like political regime type. We combine these social indicators following the shared socioeconomic pathways (SSPs) concept (Absar & Preston, 2015; O'Neill et al., 2014). By changing the values of the model inputs for the forecasts, we can conceptually “experiment” with the effects of social conditions on conflict risk as a result of the projected changes in global and local temperatures that are expected according to the Intergovernmental Panel on Climate Change (Stocker et al. 2013). In this way, our research directly addresses the question of what underlies the predicted geographical and temporal patterns.

Our theoretical starting point is that the local setting of environmental stress matters for the ability of populations to adapt. We quantify different social settings, or contexts, using projected political regime type and demographic indicators. This research represents a logical extension of prior climate change - conflict scholarship, including Burke et al. (2009) who report a null effect for the level of democracy on the risk of violence. By contrast, our results below indicate that societies and governments can effectively intervene in the face of environmental stress to mitigate conflict risk.

The quality of governance in sub-Saharan Africa, a key variable of interest in this study, is generally poor. Institutions are only partially representative and are fully autocratic in some cases, with perennially curtailed individual political and economic liberties for citizens. There are important implications of these governance restrictions for the resilience and adaptive capacity of communities in the face of climate change. For this reason, political regime type is a centrally important input variable for our five shared socioeconomic pathways (SSPs), detailed below. An environmental disaster, a devastating flood for example, can cause crops to fail, encourage urgent migration into already overpopulated areas, and decimate a national or regional economy. In a case where extreme temperatures affect crop production in only some regions of a country, repressive and autocratic regimes will be unlikely to respond adequately to the needs of affected segments of the population. This is especially true in cases where governments actively discriminate against opposition groups, often determined by ethnic affiliation, a known possibility in the clientelistic and patronage regimes that dominate our study area (Clapham, 1982; Schleussner et al., 2016). The role of governance in ameliorating or worsening resource conflicts has been highlighted since Homer-Dixon’s (1999) book on environmental security.

The premise of our study is that social institutions in a country represent moderating influences that define two pathways toward potential outcomes, one violent and the other comparatively peaceful. Where governments fail to serve the interests of the majority of citizens, social safety nets do not ease the burden of shocks to national political and economic life. Furthermore, in the private and informal

arenas of social life, the viability of economic activities and the possibilities for political expressions of grievances are limited, resulting in elevated risks of experiencing conflict (Cederman, Gleditsch & Buhaug, 2013). In contrast to the conflict-prone impacts of ineffective or discriminatory governance, countries with representative and inclusive regimes are likely to have characteristics that enable them to fare comparatively well under adverse environmental, climatological or economic conditions. There may be insurance or government aid for losses incurred by both farmers and pastoralists, for example, and this would help to sustain a household's livelihood. As a result, it is not necessary for those affected to turn to illegal or risky activities to obtain essential goods. Furthermore, good governance over the longer term can create more resilient communities, reducing the need for government assistance during times of environmental stress.

Our theoretical lens incorporates the possibility for climate and weather variability adaptation, which Adger (2006) defines as a key tool for reducing vulnerability. Dell, Jones & Olken (2014) consider social science investigations of climate change effects to be "damage function" studies where adaptation to environmental stress can occur over both short- (e.g. a farmer growing a different crop) and long-term (e.g. labor and capital migration) periods. "Government institutions and policy, including policies around public goods, innovation, and market integration, may also play important roles in the degree and nature of adaptive responses" (Dell, Jones & Olken, 2014, 772). We assume that adaptation to adverse conditions will depend on the institutional contexts present in an area. How effectively particular societies adapt to climate change is, of course, uncertain. Burke, Hsiang & Miguel (2015, 608) suggest that in East Africa between 1990 and 2009, adaptation may not have taken place, contributing to more violent conflict. Nevertheless, we believe that it is important to maintain the possibility for adaptation and situate its likelihood into our definitions of social context below.

Without effective management of agricultural production, land access, and the availability of clean water, we expect rapidly growing populations to exacerbate the social stress associated with natural resource scarcity. Our objective is to consider this simple possibility with data characterizing the political circumstances in which people compete for resources. An environmentally deterministic position would disregard our sensitivity to socioeconomic context. Parry et al. (2004) stressed that the number of people at risk of hunger in sub-Saharan Africa will rise more rapidly than that for people living in other regions. Provision of food from more productive regions can fill in some of the production deficits in sub-Saharan Africa, but consumption patterns tend to be localized in Africa and accounting for population change at sub-national scales is critically important.

To operationalize the climate-conflict nexus, we *hypothesize that conflict levels will increase along with rising temperatures if political rights deteriorate and population growth continues to rise. If sub-Saharan Africa’s political rights improve toward democratic levels and population growth moderates, conflict levels will not rise with predicted increased temperatures.* We capture these extreme cases and other possible future scenarios in our SSPs. Figure 1 presents our expectations for these relationships. While a single coefficient estimate in our statistical model could return conflict predictions (right side of the figure) based on the values for the extreme climatic conditions (left side), this would be a pathway without explanation for the observed pattern. Our approach instead leverages comparisons between scenarios in the country-level contexts (A and B in the diagram) within which social stress emerges and we attribute the possible transmission of temperature extremes into political conflict as a function of these alternative socio-political contexts. An optimistic future scenario of expanding political rights and moderate population growth is captured by Context A in the figure. Context B depicts a more pessimistic scenario where population growth rises rapidly and political rights worsen.

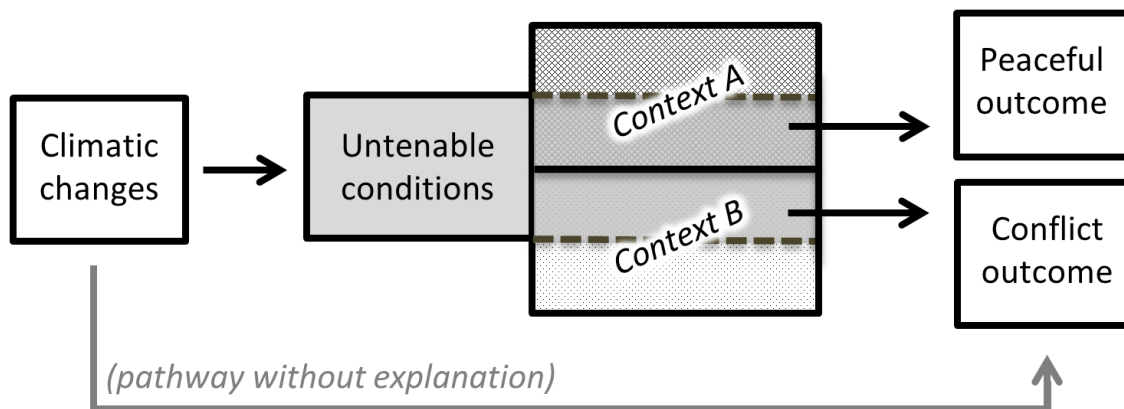


Figure 1. Schematic representation of the hypothesized relationship with socio-political contexts A and B as moderating/conditional settings (SSPs) for the effects of extreme climatological shocks.

Climate projections for Africa

Sub-Saharan Africa, particularly the Sahel and the Greater Horn of Africa, has experienced large variability in climate on inter-annual and decadal scales (Lamb & Pepler, 1992) leading to devastating

droughts, floods, and famine (Washington et al., 2006). The extremely dry Sahel of the 1970s and 1980s was associated with cooler sea-surface temperatures in the northern tropical Atlantic relative to a warmer southern Atlantic, while southern Africa's recurrent droughts seem to be associated with changes in the Indian Ocean, which has warmed more than 1° Celsius since 1950 (Hoerling, Hurrell & Eischeid, 2006).

The signals of climate change projected for Africa are emerging against this background of variability. For the continent as a whole, the mean temperature has increased since 1960 (Conway, Mould & Bewket, 2004; Kruger & Shongwe, 2004; Malhli & Wright, 2004), with more extreme hot days and nights and fewer extreme cold days and nights in southern and western Africa (New et al., 2006). Based on the reference period 1986-2005, the Intergovernmental Panel on Climate Change (Stocker et al. 2013) projects the continental mean warming trend to continue. Generally, wet areas are likely to get wetter and dry areas drier but some traditionally dry areas, such as East Africa, are projected to get wetter. The projected precipitation change for the coming decades is small compared with the magnitude of the natural internal variability of mean precipitation in Africa (Christensen et al., 2007). The sign of the projected precipitation change in both the near and long-term future has large uncertainties, with Africa and other tropical areas having the highest uncertainties in the local precipitation change (Rowell, 2011); the West African Sahel is noted for having a large spread in model projections (Roehrig et al., 2013). Uncertainties in projected precipitation change are associated with sea surface temperature (SST) changes, atmospheric and land surface processes, and the terrestrial carbon cycle. For example, in mid-century projections, warming of the Indian Ocean is associated with drying in southern Africa due to the corresponding atmospheric ascent over the ocean and subsidence over the land (Hoerling et al., 2005).

Historical and future data

To generate our forecasts of future violence, we compile historical data on factors known to influence the risk of violent conflict, and then limit these to the most important ones (our assessment of variable influences is explained below) that we can project into the future (Table I). Our forecasts do not simply extrapolate from recent data; they use covariates with known influences on conflict risk and then incorporate their future projections into our violence forecasts.

Table I. Variables used in estimating the baseline models, validation models, and future forecasts.

Variable	Data resolution		Source/notes	Grid Aggregation Method
	Temporal	Spatial		
Sociodemographic				
ACLED violent events	1980-2012 ^d	Town	ACLED v3 plus Univ. of Colorado addition	Sum
Nonviolence media reports	1980-2012 ^b	Country	Factiva	Pop-weighted mean
* Nonviolence media reports	2015-2065 ^b	Country	extrapolated from Factiva baseline	Pop-weighted mean
Population (ln)	1990-2015 ^a	2.5 minute	Gridded Population of the World, version 3 (GPWv3)	Sum, UN WPP adjusted GPW spatial distr., UN WPP urban adjusted
* Population (ln)	1950-2065 ^b	Country	UN World Population Prospects: The 2012 revision	Pop-weighted mean
* Infant mortality rate (1-yr lag)	1950-2065 ^b	Country	UN World Population Prospects: The 2012 revision	Majority pop
Political rights (1-yr lag)	1972-2012 ^b	Country	Freedom House	Majority pop
* Political rights (1-yr lag)	2015-2065 ^b	Country	Extrapolated from Freedom House baseline	Majority pop
Geographic				
* Distance to border (ln)	Constant ^e	Country	ESRI World Country Boundaries	Mean 10km sub-grid
* Capital city grid cell	Constant ^{e,f}	City	ESRI World Cities	Binary
Climate				
Precipitation (SPI6)	1949-2012 ^c	0.5°	Climate Research Unit	Mean of 1/2 degree data
Temperature (TI6)	1949-2012 ^c	0.5°	Climate Research Unit	Mean of 1/2 degree data
Temperature (TI6)	1980-2010 ^c	0.5°	Six historical simulations forced to sea surface temp.	Mean of 1/2 degree data
* Temperature (TI6)	2006-2065 ^c	0.5°	RCP 2.6, 4.5, 8.5 future coupled simulations	Mean of 1/2 degree data

^a5-year interval, ^byearly, ^cmonthly, ^ddaily

^eborders and capital city change when Eritrea (June 1993) & South Sudan (July 2011) became independent

^fcapital city changes for Cote d'Ivoire (Abidjan to Yamoussoukro, March 1983), Nigeria (Lagos to Abuja, December 1991)

* Covariate used to forecast future violence

Conflict data and media reports

To measure violence, we use an extended version of the Armed Conflict Location and Event Dataset (ACLED) which is based primarily on media reports of violence geolocated to the nearest settlement (Raleigh et al., 2010). These data code for political violence such as riots, protests, violence against civilians, and battles between rebel and government factions with a daily temporal resolution. The published ACLED data begin in 1997, and we have extended them backwards to 1980 using the published codebook and similar procedures.

Since the volume of reported violence is sensitive to the extent and depth of media coverage, we include media reports that exclude violent ones to control for the increasing volume of reports over time. These nonviolent data are derived from annual Factiva media reports for each country. This is an important component of our model since we need to ensure that we are estimating risk factors for violence over space and time, and not just capturing a technologically-driven temporal trend due to the availability of electronic event data after the mid-1990s. We project these media reports into the future by assuming they will continue to increase linearly as they have since about 1995 (see online appendix, Figures A1 and A2).

Explanatory variables

Our population metric is derived from the Center for International Earth Science Information Network (CIESIN) Gridded Population of the World (GPW), version 3 dataset (Balk & Yetman, 2013). These 2.5-minute resolution data were aggregated to our grid cells and adjusted so that the total country populations matched the United Nations (UN) World Population Prospects (WPP) 2012 report numbers (United Nations, 2014). The WPP data include consistent historical and future country-level population estimates with sub-Saharan Africa population estimates ranging from 2.2, to 2.7, to 3.2 billion in 2065 according to low, medium, and high fertility models, respectively. These latest UN data take into account the upwardly revised fertility rates for Africa (Gerland et al., 2014). By standardizing the spatially disaggregated GPW to the WPP data, we can use WPP forecasts without introducing any discontinuities into the population figures.

For future population data, we use the GPW spatial distribution (2.5-minute resolution) to allocate the UN WPP country-level projections for the low, medium, and high fertility scenarios. The spatial allocation is not uniformly proportional into the future since we modify the estimate to take into account UN WPP urbanization projections, thereby assuring that urban areas grow faster than rural areas and thus receive a greater share of future population growth. Urban population growth is added to current urban areas, and rural population totals are matched to rural grid cells. These fine-resolution allocations are then aggregated to our 1° unit of analysis.

Socioeconomic well-being is important because locations in poor areas typically have higher levels of violence. To capture this influence as others have (Theisen, Holtermann & Buhaug, 2011), we use the infant mortality rate as a proxy for well-being. In contrast to indices of wealth, IMR is easier to measure and more reliable; therefore, the UN WPP historical and future data do not suffer from many missing values (United Nations, 2014). For future analyses, it may be possible to use nighttime lights as a surrogate for local wealth estimates as have Weidmann and Schutte (2017), though the temporal range of the nighttime lights data limits their usefulness.

We also include a measure of governance that captures the political rights for each country. These political rights data from Freedom House (7 = least free, 1 = most free) measure the effects of formal political institutions (Freedom House, 2013). Both the IMR and political rights variables are lagged one year to mitigate endogeneity with the outcome measure, conflict events. To estimate governance in the future, we consider three scenarios (see SSP definitions below for combination with other data). In the pessimistic outlook, we assume that African political rights will decline to the levels of the 1980s. In the

optimistic scenario, political rights for every country improve (the scores improve towards a value of 1) by following the linear trend for all of sub-Saharan Africa, 1980-2012. Our middle-of-the-road scenario holds political rights constant for each country (online appendix Figures A3 and A4).

Since borders are areas that frequently experience higher levels of conflict as we have shown in our previous work (O'Loughlin et al., 2012; O'Loughlin, Linke & Witmer, 2014), we include a distance to border metric by calculating the mean distance to international borders for a 10-km subgrid, and then aggregate these to our 1° grid. To further examine geographic effects, we add a binary variable for the grid cell that contains the capital city of each country. This factor is especially important for capturing political violence targeting the incumbent regime since capital cities tend to be the sites of major violent protests and rioting, and these events form part of the ACLED data.

Climate data

Historical climate data are derived from the 0.5° monthly Climate Research Unit (CRU) TS3.21 dataset from the University of East Anglia (Harris et al., 2014). The CRU data cover land areas only and include the number of stations used to interpolate each grid value, which allows the reliability of the values to be determined objectively. From these data, we calculate temperature and precipitation anomaly indices by comparing the most recent 6 months of data with the prior 30-year climatology for those same 6 months and grid locations. Our use of a running 30-year climatology instead of a fixed climatology allows for the possibility that societies will adapt to changing conditions. For rainfall, this is the Standardized Precipitation Index, SPI6, normalized using an incomplete gamma distribution where values near 0 indicate normal precipitation and -1 indicates that the last 6 months were 1 standard deviation drier than usual (McKee, Doesken & Kleist, 1993). The temperature index, TI6, is calculated in a similar manner (though using a standard normal distribution) with positive values indicating hotter than usual temperatures. This normalization process allows us to directly evaluate the effects of precipitation and temperature anomalies across grid cells with different climates.

We use climate models in two ways, to validate our violence model and to project future climate conditions. For both the historical and future climate forecasts, we use simulated climates generated from the Community Earth System Model (CESM; Hurrell et al., 2013). Though there are other climate models that could be used (Burke et al., 2015), CESM uses a wide range of future emissions possibilities that covers the range of future climate scenarios, and is well validated for the present climate. To validate our statistical model of violence, we estimate it using the simulated historical climate data generated by running the atmosphere and land model with monthly mean observed sea surface

temperature (SST). This ensures that large-scale climate patterns from surrounding oceans are used to force the model, and for regions where the modes of SST variability are important, this reduces climate variability and variance from historical observations. Land temperatures and the atmospheric hydrologic cycle (including precipitation) are allowed to freely evolve. The benefits can clearly be seen in the temperature anomaly figure where present day (1980-2010) model simulations closely track observed regional SST anomalies (Figure 2). Six different simulations are used: they vary only in an initial round-off level perturbation to the surface pressure field to set the ‘weather’ states on different trajectories. The resulting climate anomalies are similar, and the spread relates to the remaining unforced internal variability in the model.

Sub-Saharan Africa Temperature Anomaly

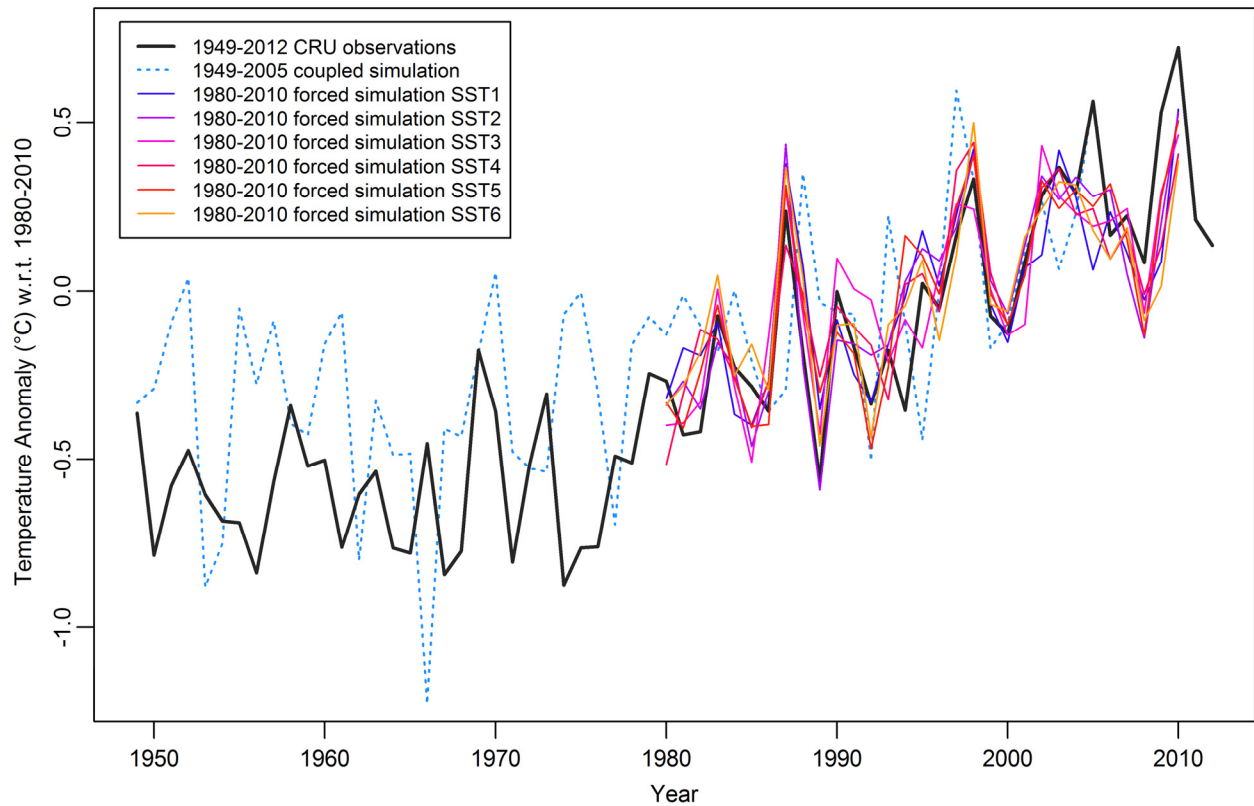


Figure 2. Sub-Saharan Africa historical temperature anomalies 1949-2012 for CRU data, a coupled simulation, and six specified sea surface temperature (SST) simulations.

We also run future projections with the CESM. Future simulations are conducted with the simulated atmosphere model coupled to a dynamic ocean model. Simulations use several different anthropogenic

emissions scenarios or Representative Concentration Pathways (RCPs) that specify the climate forcings, in particular, CO₂ levels (Van Vuuren et al., 2011). In this article, we use a range of scenarios from a ‘high’ or baseline emissions case with 8.5 Wm⁻² of radiative forcing by 2100 (RCP8.5), a ‘moderate’ case assuming 4.5 Wm⁻² of forcing (RCP4.5), and a ‘low’ emissions case of 2.6 Wm⁻² forcing (RCP2.6). The scenarios are based on integrated assessment models that include assumptions about population projections, technology improvements and possible limits on emissions of greenhouse gases due to national policies. RCP8.5 features a ‘wealthier’ world with higher population growth than the other scenarios, but all three are deemed plausible (Stocker et al., 2013).

Future climate simulations use the fully coupled version of CESM with the three different scenarios (RCP8.5, RCP4.5 and RCP2.6). We extract monthly mean physical climate statistics for the study area from the model outputs and use it as a ‘synthetic’ set of climate data to calculate future precipitation and temperature anomalies for input to the statistical models. As with the historical fixed SST scenarios, multiple simulations with small perturbations between them are used to sample the possible internal variability in the model for each scenario. To maintain internal consistency, future precipitation and temperature anomalies are calculated against a rolling 30-year climatology from a corresponding coupled historical simulation, 1949-2005 (Figure 2). This coupled simulation is forced by greenhouse gases and observed natural forcing such as volcanic eruptions. It will not reproduce the exact timing of internal modes of variability (such as the El Niño-Southern Oscillation) as in the uncoupled simulations with fixed sea surface temperatures.

Baseline models

Our spatio-temporal units of analysis are the 2,062 1° x 1° grid cells overlaid on 42 countries of sub-Saharan Africa for each month (see Figure 5 for their spatial distribution). The 33 years of data yield 816,552 grid-month observations used to calibrate our statistical models. Our approach is to generate estimates for our baseline models, and then use our projected covariates to generate future forecasts. We use a Poisson¹ multilevel model with country-level random effects and log link function of the form:

$$y_{ij} \sim \text{Poisson}(\lambda_{ij})$$

$$\log(\lambda_{ij}) = \beta_{0c} + \beta_1 X_{ij} + \varepsilon_{ij}$$

¹ We also tested alternate functional forms (linear, logit, and negative binomial) with little change to the coefficients.

where β_{0c} are the country-level intercepts, β_1 are the coefficients (fixed effects) for the grid-month predictors X_{ij} , and ε_{ij} captures the remaining unexplained error. Our motivation for using a multilevel model is to capture the nested structure of the relationship we investigate, allowing spatially disaggregated processes among units of analysis to take place within broader social settings. We also prefer our random effects specification to fixed effects, due to the large number of non-conflict observations that do not contribute to the statistical analysis when using a fixed effects specification (Beck & Katz, 2001).

Since our models have a large number of observations, the usual standard errors are underestimated. Robust clustered standard errors are frequently used to address this situation, but since they are typically not used with multilevel models, we empirically calculate the standard errors from 200 bootstrapped coefficient estimates. Given the multilevel structure of our data, we employ a hierarchical resampling approach that reflects the data generating process as closely as possible (Davison & Hinkley, 1997). In particular, we resample the observed data with replacement for the country level first, then grid cell level, and then monthly level.

Our initial baseline model, Model 1, includes the full suite of sociodemographic, geographic, and climatic predictor variables (Table II). Statistical significance from the bootstrap standard errors indicate that the coefficients for precipitation anomalies, well-being, and distance to border do not differ from 0. Model diagnostics for area under the receiver operator curve (AUC-ROC), the precision-recall curve (AUC-PR) and the Brier score are calculated for our count data by truncating predicted values above 1.

Table II. Poisson random effects multilevel models (MLM-RE), 1980-2012.

Fixed part	Model 1		Model 2, no SPI6	
	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>
Constant	-16.628	2.089 **	-16.638	2.205 **
Precipitation (SPI6)	0.028	0.032		
Temperature (TI6)	0.108	0.042 *	0.104	0.041 *
Population (ln)	0.820	0.130 **	0.819	0.140 **
Well-being (IMR lag)	0.382	0.267	0.381	0.282
Political rights (lag)	0.337	0.123 **	0.337	0.132 *
Capital city grid cell	1.004	0.354 **	1.003	0.379 **
Distance to border (ln)	-0.147	0.118	-0.146	0.113
Non-violence media trend (ln)	0.379	0.082 **	0.381	0.082 **
Random part				
Country level σ^2	0.8824		0.8828	
Model diagnostics				
AUC-ROC	0.8454		0.8454	
AUC-PR	0.1582		0.1581	
Brier score	0.0379		0.0379	

Significance codes: ** $p < 0.01$, * $p < 0.05$. σ^2 is variance.

Number of observations (grid-months) = 816,552.

Random effect intercepts not reported for 42 countries.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

We explore the contribution of each variable further by plotting the predictive power contribution of each variable as captured by the AUC-ROC against the z value from the bootstrapped standard error (Figure 3). This step provides additional information by which to evaluate the importance of each variable for violent conflict. To avoid introducing excess random noise into our future forecasts, we drop precipitation (SPI6) from Model 1, and re-estimate the coefficients in a trimmed version, Model 2 (Table II). Although this model does not have an explicit precipitation metric, radiative and sensible heat fluxes are products of temperature and guide the local precipitation response over land (Muller and O’Gorman, 2011). Material well-being (IMR) and distance to border are retained since they contribute to the overall predictive power of the model (Figure 3). Our finding for precipitation runs counter to some expectations in the literature (e.g. Miguel, Satyanath & Sergenti, 2004), but there are several possible explanations for the alternative conclusions. Miguel, Satyanath & Sergenti (2004) study civil war at a country level, where a year must count at least 1,000 deaths in confrontations between a government and cohesive rebel organization. Instead, we focus our attention on the kind of low-level deadly violence that is more common in sub-Saharan Africa recently. At a local level, populations often migrate in a time of drought; Miguel, Satyanath & Sergenti (2004) cannot capture these sub-national level reactions to precipitation variability.

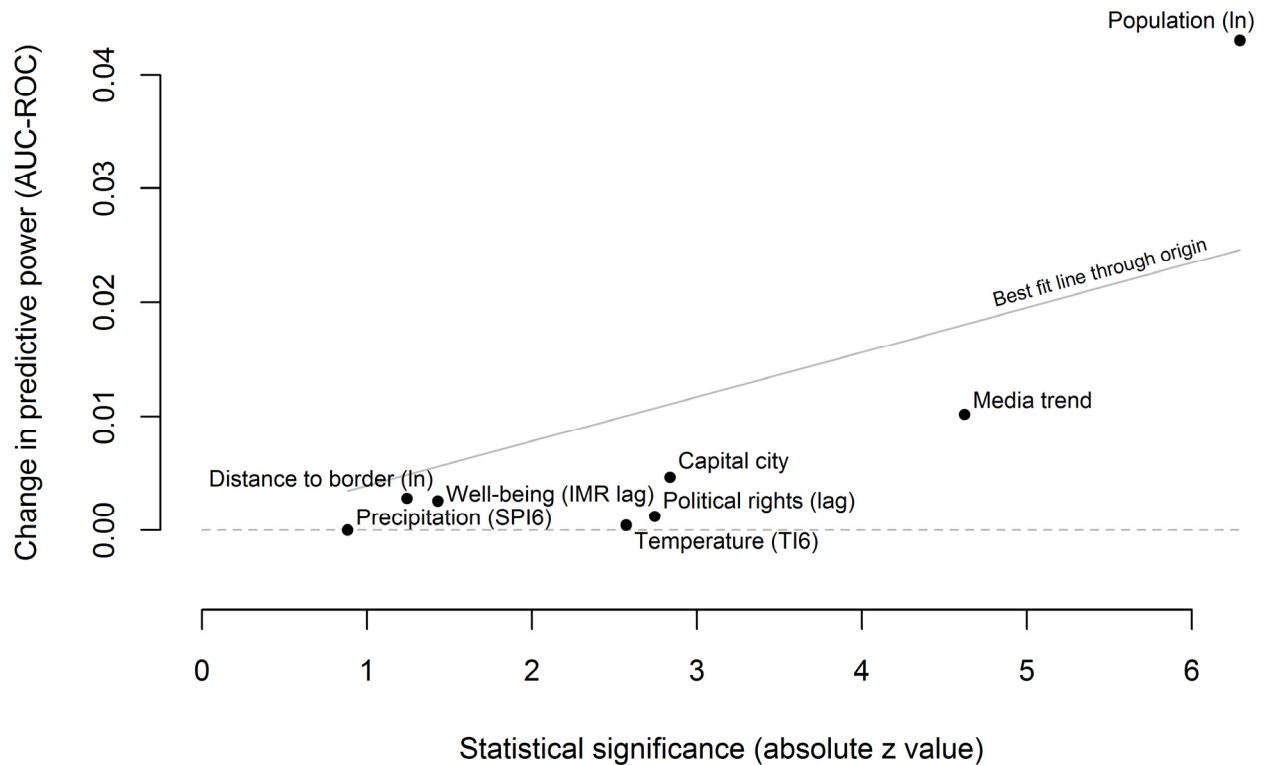


Figure 3. In-sample predictive power plot for Model 1.

To increase confidence in our modeling decisions, we estimate a set of alternate models and describe the results in the online appendix. We change the measure of violence by predicting the ACLED sub-types representing violence against civilians, riots/protests, and battle events (Table A1). In addition to these sub-types, we replace the ACLED dataset by testing the model specification using the Uppsala Conflict Data Program Georeferenced Event Data (UCDP-GED) data (Croicu & Sundberg, 2015; Sundberg & Melander, 2013). Table A2 shows results for this model and for two variants that drop the non-violence media reports and exchange the political rights metric for a polity score (Polity IV Project, 2014).

We also estimate models with varying drought representations by explicitly combining temperature and precipitation. Table A3 shows these results for a hot and dry measure (a simple difference: TI6 – SPI6) and also for a more sophisticated 6-month Standardized Precipitation Evapotranspiration Index (SPEI6) that combines the precipitation and temperature record with the latitude for each grid cell. Finally, we present the variants of the main Table II models that calculate the temperature and precipitation anomalies using the long-term climatological record from 1949-2012 instead of the 30-year rolling climatologies (Table A4). These results in the online appendix underline and support our main findings from Table II (full country-level random effects are reported in Table A5 and Figure A5). The

online appendix also presents an out-of-sample model validation (Table A6) and models using the six historical climate simulations (Table A7).

Violence forecasts for sub-Saharan Africa, 2015-2065

To forecast violence, we use coefficients from Model 2 estimated from the full duration of observed data, 1981-2012. We focus our results on five future scenarios (Absar & Preston, 2015; O'Neill et al., 2014) that allow population growth, political rights, and climate projections to vary in line with what we might expect for each of the shared socioeconomic pathways (SSPs). Explicitly modeling future variation in political rights within the SSPs is reasonable since democratizing trends and reversals in political and civil rights are well-documented for sub-Saharan Africa (Lindberg, 2005):

SSP1: 'Optimistic' future. This is a best-case scenario where population growth is low and political rights improve. Temperature simulations use RCP2.6.

SSP2: 'Middle-of-the-road' future. This is a mid-course scenario with medium population growth and constant political rights. Temperature simulations use RCP4.5.

SSP3: 'Pessimistic' future. This is a worst-case scenario where high population growth is combined with a decline in political rights. Temperature simulations use RCP8.5.

SSP4: 'Inequality' future. Globally population growth varies but remains high for poorer areas such as sub-Saharan Africa. Political rights remain constant and temperature simulations use RCP4.5.

SSP5: 'Contrasting growths' future. Low population growth and political rights improvements are combined in this scenario of best social outcomes and worst climate changes (RCP8.5) due to fossil-fueled development.

To forecast violence for each of the scenarios, we use a simulation approach to generate 200 futures. For the fixed effects coefficients, we use the 200 bootstrap coefficients used to generate the standard errors of the estimates. Our bootstrap resampling approach means that coefficients for some of the country level random effects are missing, so we simulate 200 random effect coefficients and pair them with the bootstrap fixed effects. Then, for each scenario, we use the specified population and political rights projections, and use climate outputs from two simulations (same RCP), applying 100 simulated coefficients to each.

Though we simulate the intercept values for the country-level random effects in the forecasts, the means for each country remain constant. While this assumption can be justified in terms of risk factors that are generally stable over time (e.g. terrain, soil quality, access to strategic points), it does not capture unobserved country changes such as infrastructure development and other policy initiatives.

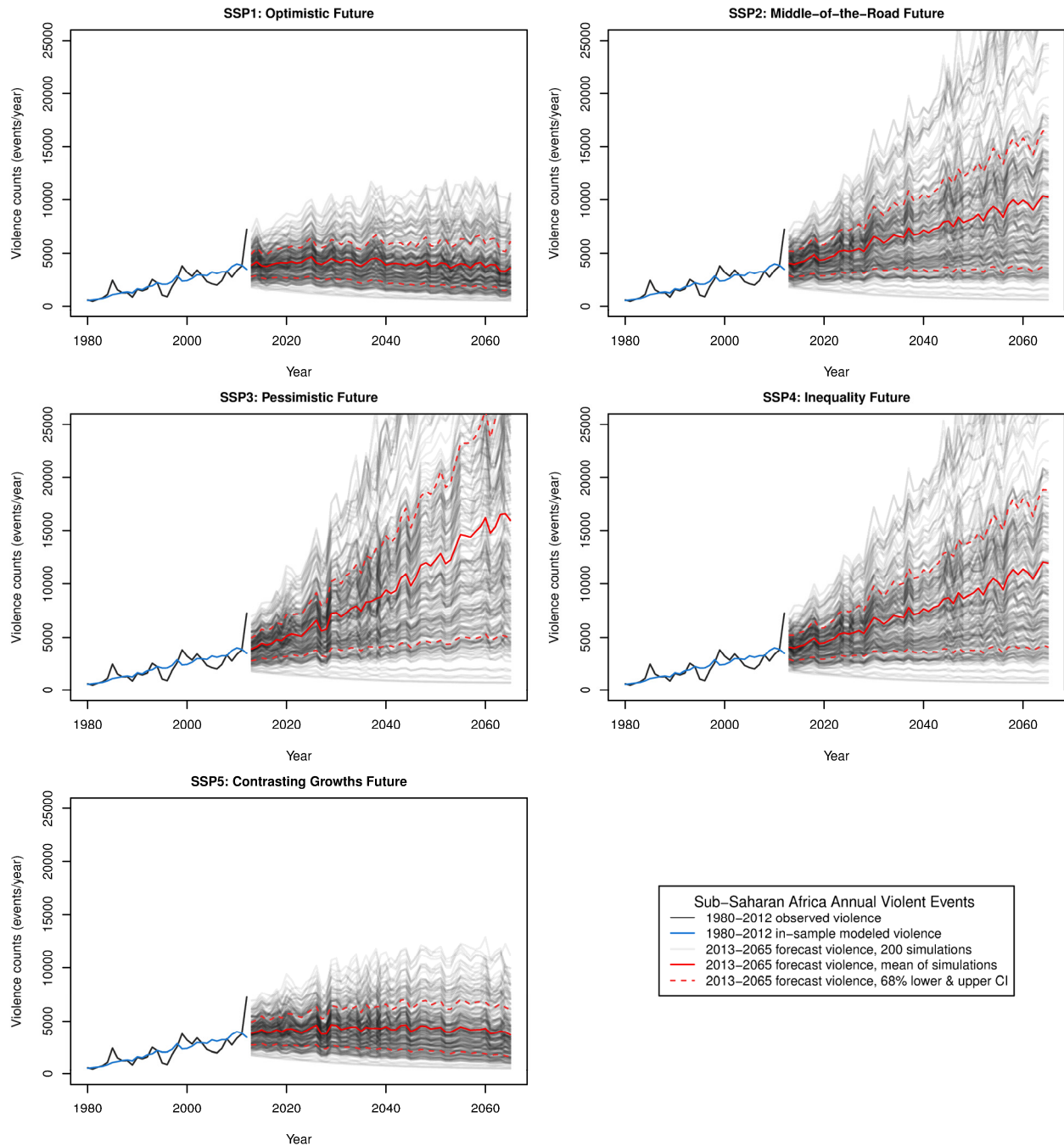


Figure 4. Forecast annual violent event counts for all of sub-Saharan Africa through 2065 for each of the defined shared socioeconomic pathways (SSPs).

Figure 4 shows the results for each of the five scenarios by plotting total annual forecasted violence in sub-Saharan Africa. Based on our assumptions, shared socioeconomic pathways 1-3 show the complete range for forecast violence, from decline in the optimistic scenario, to moderate increases for the middle-of-the-road scenario, to a tripling of violence in the pessimistic case. SSP1 and SSP5 forecast the least amount of violence and are characterized by improving political rights and relatively low population growth. This general downward pattern coincident with improved political rights is a similar result to Hegre et al. (2013), who find a decreasing trend for intra-state civil war in a similar scenario.

Our forecasts also identify conditions under which violence could rise with increasing temperatures and pessimistic socio-political futures. SSP3 and SSP4 forecast a dramatic rise (Figure 4) in conflict though 2065 under scenarios where political rights in sub-Saharan Africa remain poor. In our formulation, the only difference between SSP1 and SSP5, and SSP3 and SSP4 are the climate projections. In our model forecasts, future violence is sensitive to changes in political rights and governance, and relatively insensitive to temperature anomalies. This conclusion stands in dramatic contrast to some hyperbolic scenarios about violent social reactions to climate change (White House, 2015).

The advantage of our comparative modeling strategy across SSPs is that it considers the capacity of societies and institutions to manage the stress that can emerge with temperature increases, including droughts and increasing variability in water access. The degree to which people trust in social adaptation and anticipate progress in political rights through the mechanisms of representative governance affects their responses, including conflict options.

The fine spatial (1° grids) and temporal (monthly) detail of our data allow us to map our forecasts. The forecast maps use the estimated (not bootstrap or simulated) model coefficients and the mean temperature anomalies for two simulated future climates (same RCP) to forecast violence. Figure 5 maps total violence for the middle-of-the-road scenario in 10-year periods to minimize inter-annual variation. As expected, the future spatial distribution reflects recent levels of violence as predicted by population distributions, distance to borders, and locations of capital cities. Country-level influences are also visible (e.g. for Niger), with such responses driven by a country's political rights.

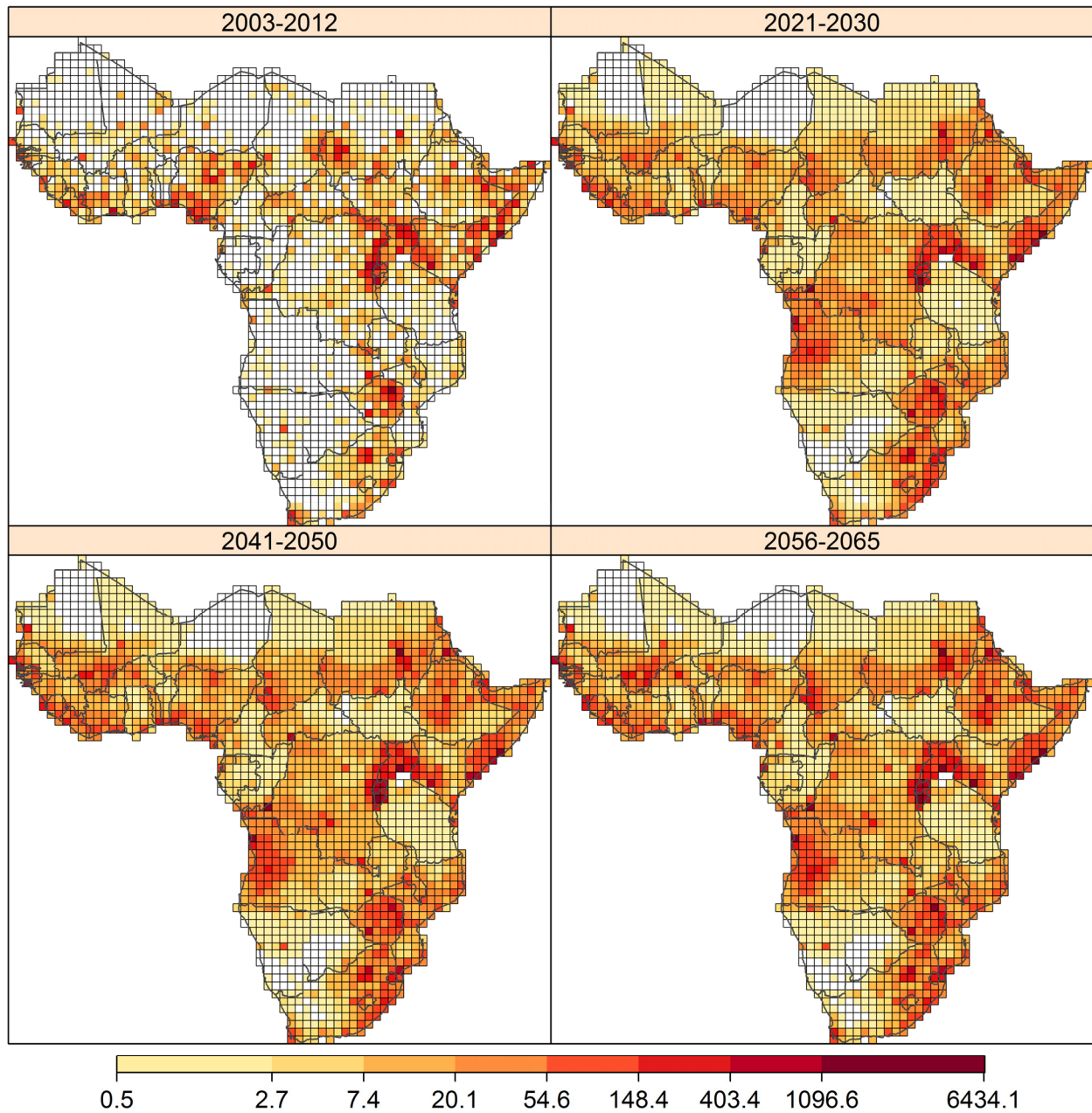


Figure 5. Observed and forecast violence for SSP2, middle-of-the-road future.

To examine the numerical distribution of the decadal violence in Figure 5, we use a grouped histogram where each bar tone is associated with a given decade (Figure 6). Note that the event count bin sizes vary to accommodate the skewed distributions. The generally increasing violence forecast with SSP2 (Figure 4) is visible in the histogram, with counts declining for grid cells with fewer than 1 violent event, and other bar heights generally rising, especially for the most violent grid-months forecast to

experience over 150 events. Indeed, much of the increase in total violence is being driven by grid cells that experience high levels of violence. From the 2040s to the 2055s periods, 809 grid cells were forecast to decline in violence (on average -0.7 events), while the remaining 1,253 locations were forecast to increase in violence (on average 9.6 events). This suggests that even under a scenario where overall violence increases, many places (though generally less populated) are expected to experience less violence.

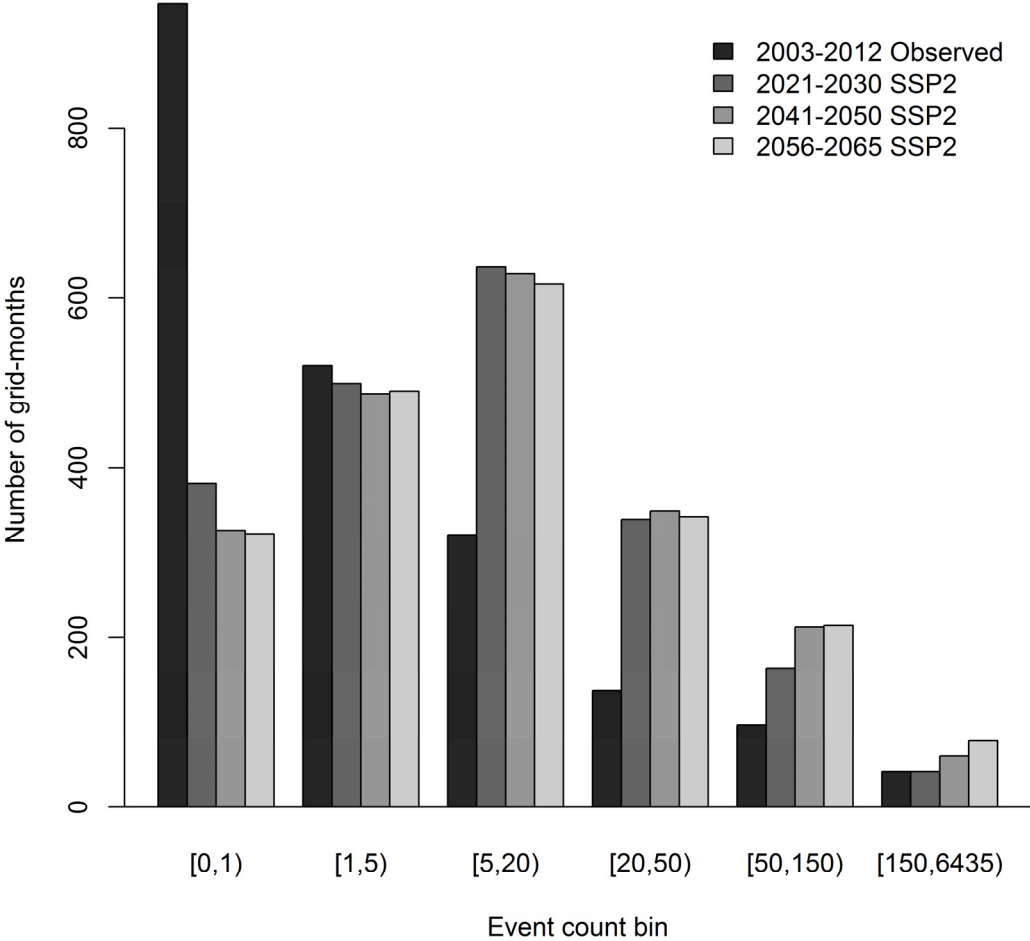


Figure 6. Histogram of violent event counts for the latest observed decade and three future decades as forecast by SSP2, the middle-of-the road scenario.

To display more clearly the change from recent conflict levels to forecasted levels, we map the ratio of violence in the 2056-2065 decade to the period 2003-2012 (Figure 7). Blank grid cells indicate no violence was observed from 2003-2012. Increases in violence are forecast for all scenarios in countries such as Sudan, Ethiopia, and Angola, driven in part by continued low political rights and high population

growth. For futures marked by improving political rights (SSP1 and SSP5), areas of decreasing violence are clearly visible in South Sudan, Nigeria, and the Ethiopia-Somalia border region. Even for SSP2, despite an overall increasing trend in violence (Figure 4), there are large areas where violence is expected to decrease (northern Mali, South Sudan, northern Kenya and Namibia). In the most pessimistic scenario, SSP3, almost all areas will experience an increase in conflict, including relatively peaceful countries such as Tanzania. An exception for SSP3 is South Sudan, which is forecast to experience a reduction in violence. This is a result of the large negative coefficient on the South Sudan random effect (Table A5 and Figure A5), likely caused by few observations in the dataset for this young country. Comparing SSP3 to SSP1 for West Africa reveals a dramatically more violent future when political rights deteriorate. Such dramatic spatial variation is all too often lost in country-scale analyses.

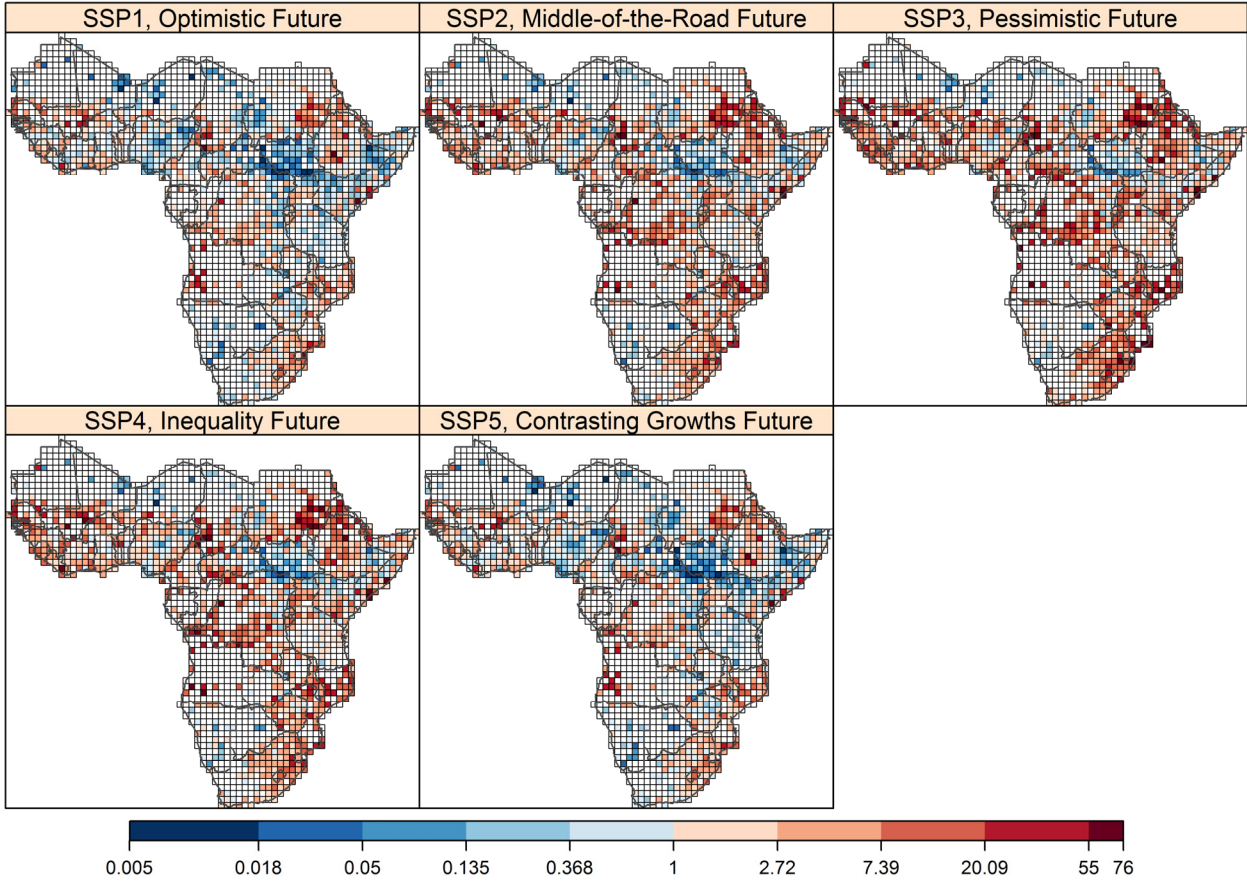


Figure 7. Decadal ratio of 2056-2065 forecast to 2003-2012 observed violence for each scenario.

These SSP-driven forecasts are instructive for comparing plausible alternative futures, but do not allow us to easily isolate the relative contributions from any one factor since most scenarios differ in two or more ways. To isolate the effects of political rights, population, and temperature anomalies, we use SSP1, the optimistic future, as a baseline future and modify it in turn for each of the factors. The three maps in Figure 8 show the differences when governance deteriorates towards autocracy compared to the improving trend of SSP1, when population growth is marked by high fertility compared to the low fertility of SSP1, and when temperatures increase from the RCP 2.6 of SSP1 to the expected higher temperatures of RCP 8.5. These maps show stark differences, with changes in the nature of governance clearly having the greatest impact on future violence levels. Isolating these governance factors also helps to explain the temporal trends (Figure 4) and spatial variation (Figure 7) in forecast violence between the SSPs, with changes in political rights driving much of the projections.

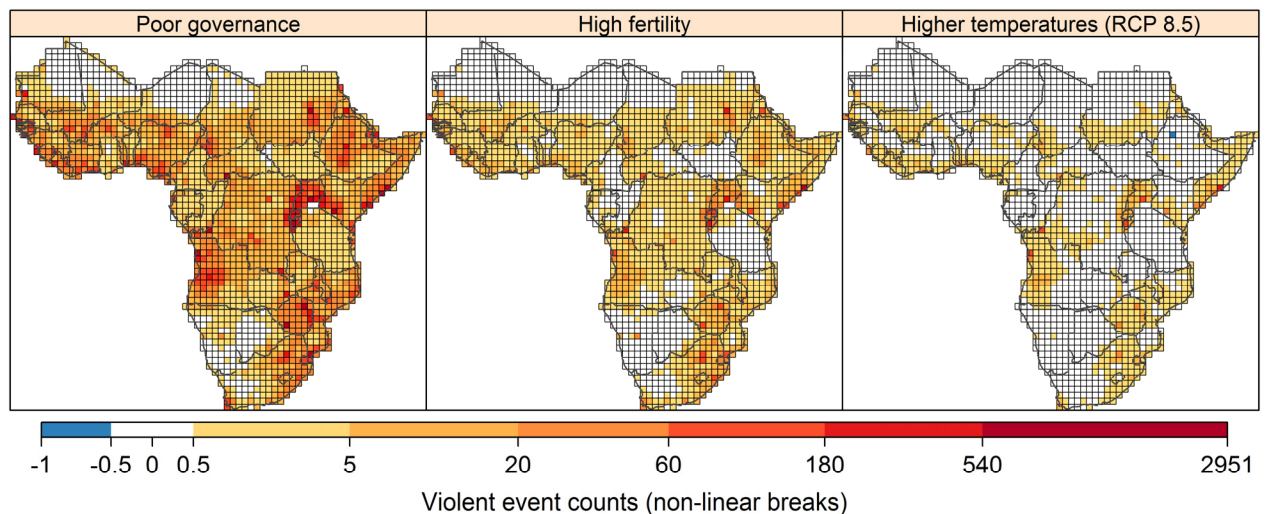


Figure 8. Change in forecast violence for the period 2056-2065 using SSP1 as the baseline forecast and comparing against poor governance, high fertility, and higher temperatures.

To evaluate the extent to which these differences in forecasts are meaningful, we use a Bayesian estimation method to compare the forecast levels of violence for the same three sets of scenarios shown in Figure 8. Given a set of two input datasets, their means and standard deviations are used to calculate 100,000 credible parameter-value combinations using Markov chain Monte Carlo (MCMC) simulation (Kruschke, 2013). The resulting posterior distributions can then be used to evaluate if the difference in means of the two input datasets are credibly different, i.e. do not overlap with 0. In our

case, each input dataset consists of the 200 forecast violence simulations (grey lines in Figure 4). To reduce inter-annual variation, we calculate the 2056-2065 mean for each forecast.

Figure 9 shows that, compared to the baseline optimistic SSP1 scenario, futures marked by poor governance or high fertility will have credibly higher levels of violence, whereas the future with higher temperatures (RCP 8.5) differs little from a future with less warming (RCP 2.6). If we wished to forecast violence to 2100, the temperature effect would likely be more pronounced, as RCP 8.5 projects the latter half of the century to warm significantly (Stocker et al., 2013). See the online appendix Figure A6 for plots that show which years differ from the baseline SSP1 scenario using this Bayesian technique and an alternate *t* test.

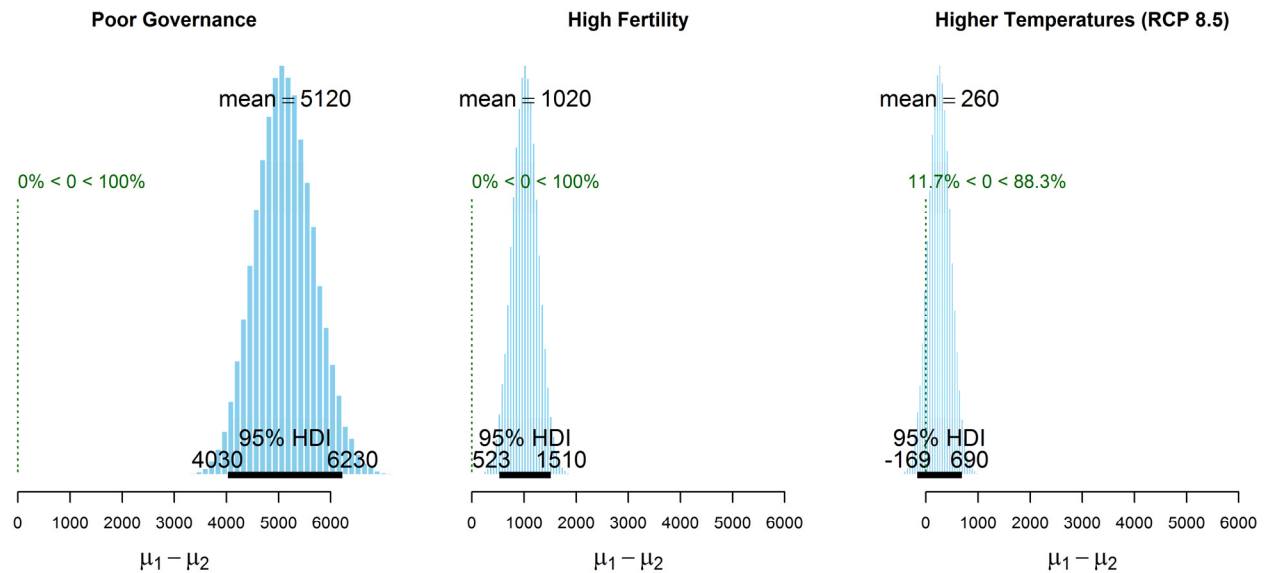


Figure 9. Posterior distributions of the difference in mean forecast violence for the period 2056-2065 using SSP1 as the baseline forecast and comparing against poor governance, high fertility, and higher temperatures. The black bar at the bottom represents the highest density interval (HDI) and shows where 95% of the credible values fall.

Conclusions

In this paper, we use spatially and temporally disaggregated local violent event data and simulated climatological data based on varying Representative Concentration Pathways (RCPs as used by Stocker

et al., 2013) to forecast future levels of violence under plausible alternative social, demographic, and political scenarios. We find statistically significant relationships between the amount of violent conflict and political rights, population size, and rising temperatures. We find no relationship with precipitation anomalies, but our analysis differs in important ways, especially in the level of disaggregation, from earlier work reporting this association. Our model validation using simulated historical climate data revealed that the temperature relationship is specific to the actual observed data compared to the simulated data that, though they match the overall trend, allow for localized variation across time and space.

For the forecasts of violence, under pessimistic future scenarios of unchanging political rights from current levels (SSP3 and SSP4), rising temperatures and increasing population exacerbate levels of violence. If political rights and governance improve (SSP1 and SSP5), then future levels of conflict are likely to remain stable or even decline, despite increases in temperatures and population. This result should caution against exaggerated claims of a violent future that are based on inappropriate data, incorrect geographic specifications, and an under-appreciation of the key role of good governance in the dampening of conflict through a fairer distribution of scarce resources (Homer-Dixon, 1999).

Some research studying the conflict effects of climate change alludes to rising risks of violence coincident with and connected to global warming. In fact, this is a key emphasis where authors assert the practical importance of their research for the general public and policy options. Our article has quantified and formalized the study of these associations with the best available predictions of temperature and precipitation into the future. As others have done for the study of democracy and conflict (Hegre et al., 2013), we extend the study of violent conflict to forecast violence using plausible future sociopolitical scenarios (SSPs). As a contribution to the conflict forecasting literature, to our knowledge, our spatially disaggregated approach is the first of its kind.

In terms of policy implications, the results of our study are sufficiently persuasive to maintain and strengthen efforts aimed at improving political rights, political freedom, and good governance. For decision makers, policies aimed at improving governance are likely to be effective in achieving the goal of reduction in violent conflict even if efforts to reduce greenhouse gas emissions are tardy or timid. Our results indicate that one major pathway for human intervention – institutional – can in fact have powerful effects on future levels of conflict given the expected changes in climatological and environmental conditions.

Replication data

The online appendix, dataset, and R replication files for the empirical analysis in this article are available at <http://www.prio.org/jpr/datasets>.

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FRANK D W WITMER, b. 1975, PhD in Geography (University of Colorado, Boulder, 2007); Assistant Professor, University of Alaska Anchorage (2014—); current main interests: remote sensing and spatial statistical analysis of violent conflict, land-use/land-cover change.

ANDREW M LINKE, b. 1983, PhD in Geography (University of Colorado, Boulder, 2013); Assistant Professor, Department of Geography, University of Utah (2015—); current main interests: political violence and conflict, political geography, and spatial statistical analysis.

JOHN O'LOUGHLIN, b. 1948, PhD in Geography (Pennsylvania State University, 1973); College Professor of Distinction and Professor of Geography, University of Colorado, Boulder (1988—); current main interests: climate changes and social outcomes in sub-Saharan Africa, public opinion surveys, conflicts in the former Soviet Union, political geography.

ANDREW GETTELMAN, b. 1970, PhD in Atmospheric Sciences (University of Washington, Seattle, 1999); Scientist III, National Center for Atmospheric Research, Boulder, CO (2010—); current main interests: clouds and climate, climate simulation.

ARLENE G LAING, PhD in Meteorology (Pennsylvania State University, 1996); Research Scientist, Cooperative Institute for Research in the Atmosphere, Colorado State University; current main interests: extreme precipitation, climate variability, convection, tropical meteorology, and numerical weather prediction.

Sub-national violent conflict forecasts for sub-Saharan Africa, 2015-2065, using climate-sensitive models

Frank D W Witmer, Andrew M Linke, John O'Loughlin, Andrew Gettelman, Arlene Laing

Journal of Peace Research

Online Appendix

Our online appendix provides additional information for the nonviolence media reports and political rights scores, two key variables in our analysis. We also include a set of additional models that serve as robustness checks for the models presented in the main text. For completeness, we report the full country-level random effects results for our main models. We also show results from an out-of-sample model validation and models using the six historical climate simulations. Lastly, we expand on the evaluation of conflict forecast differences by reporting yearly results using both the Bayesian difference in means reported in the main text and an additional t test.

Nonviolence media reports

We use nonviolence media reports to adjust for any bias in the reporting of violence associated with increasing media coverage over time. These data were compiled using repeated Factiva searches for each country with violence terms excluded, i.e. "NOT conflict NOT attack NOT violence NOT strike NOT battle NOT war NOT army NOT rebel NOT explosion" (*citation omitted during review*). Figure A1 shows these observed count data from Factiva searches, 1980-2012, and a linear projection into the future. The linear fit is calculated from a subset of the observed data, considering values only since 1995. This trend is then extended into the future and applied to each country individually from their respective 2012 starting point. Figure A2 shows the natural log transformed version of these annual sub-Saharan data.

Media Trend for Sub-Saharan Africa

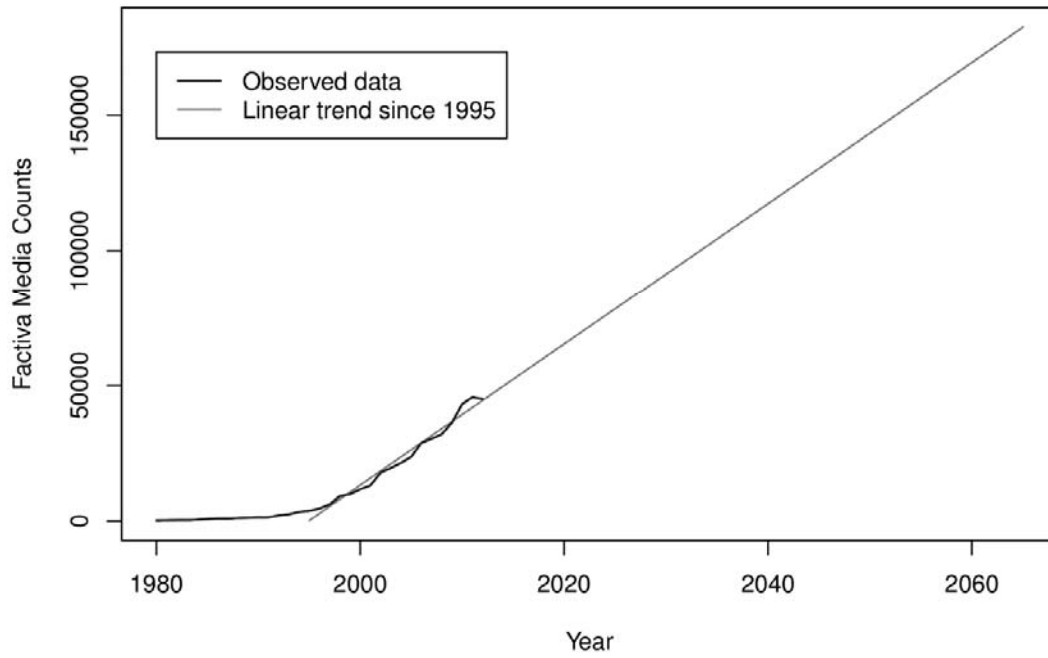


Figure A1. Observed and projected non-violence media reports. Units are total annual reports for all of sub-Saharan Africa.

Media Trend (ln) for Sub-Saharan Africa

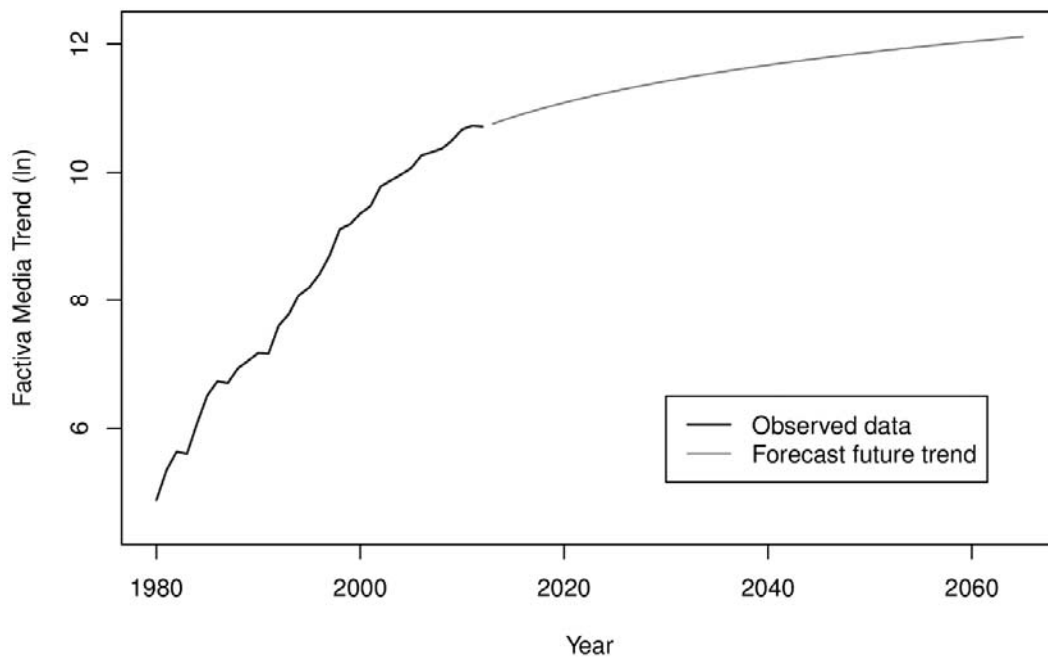


Figure A2. Natural log transformed observed and projected non-violence media reports for all of sub-Saharan Africa.

Political rights projections

Figure A3 shows the observed mean political rights scores from the Freedom House data for 1980-2012, and for three future projections. The pessimistic future scenario returns governance values to those experienced during the 1980s, while the optimistic future continues the current linear trend of the data into the future. The given projected trend is then applied to each country starting at the observed 2012 score. The best score cannot be below 1, the highest political rights freedom value. The resulting spatial distribution for projected values is shown in figure A4.

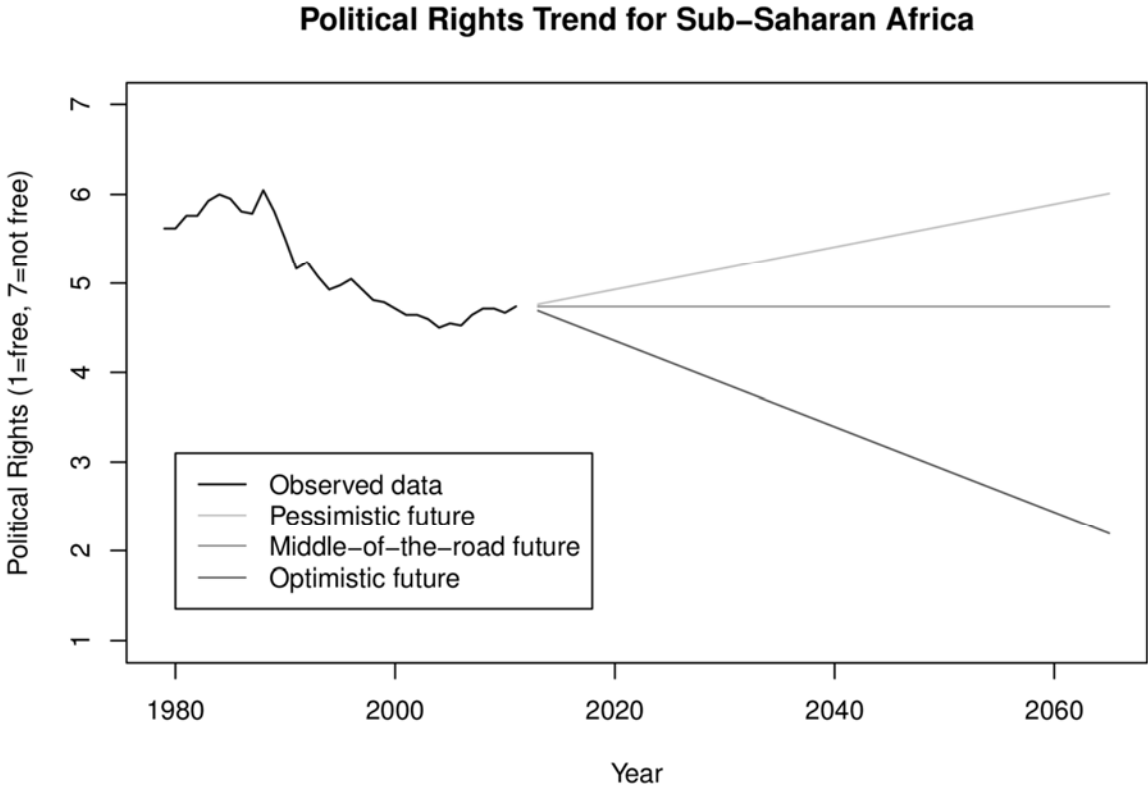


Figure A3. Mean annual political rights score (1 = free, 7 = not free) for all of sub-Saharan Africa. Observed data for 1980-2012. Gray lines show projected linear trends for future years.

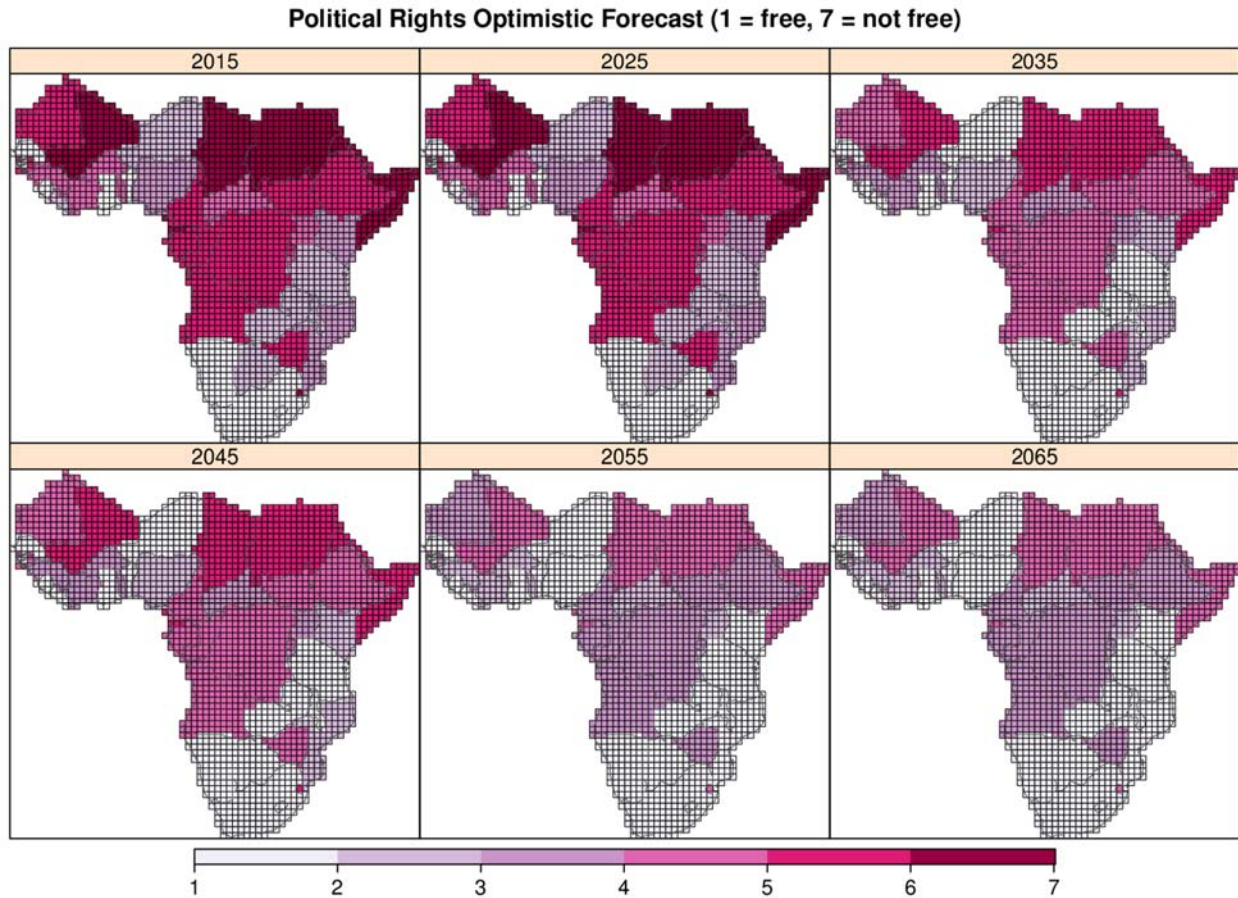


Figure A4. Spatial distribution of gridded country-scale political rights (Freedom House) scores (1 = free, 7 = not free) from the observed data in 2012 projected to select future years.

Additional models

Before presenting the additional models, a note about the reported model diagnostics. The receiver operator curve (ROC) is formed by plotting the true positive rate against the false positive rate, while the precision-recall (PR) curve plots the positive predictive value against the true positive rate. For the AUC-ROC, values closer to 1 indicate better predictive power, while AUC-PR values closer to 0 are desirable. For a further discussion of these measures, see Ward & Beger (2015). The Brier score is the mean squared difference between the predicted and observed violence, and scores closer to 0 indicate a better predictive model.

The ACLED violent event data are coded to distinguish different types of events, violence against civilians, riots/protests, and battle events. Table A1 shows the results from the initial model specification

(Table 2, Model 1) with the outcome variable replaced with each of these three violence sub-types. Model estimates are generally similar, though as expected they exhibit variation for some variables such as temperature anomalies (only significant for battle events), political rights (not significant for riots/protests) and well-being (only significant for riots/protests).

Table A1. ACLED models by violence sub-type: violence against civilians, riot/protests, and battles events.

	Violence against civilians		Riot/protest events		Battle events	
	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>
Fixed part						
Constant	-18.949	2.804 **	-20.498	1.835 **	-16.522	2.236 **
Precipitation (SPI6)	0.036	0.052	-0.054	0.047	0.048	0.036
Temperature (TI6)	0.109	0.067	0.002	0.043	0.129	0.059 *
Population (ln)	0.852	0.184 **	1.171	0.140 **	0.708	0.124 **
Well-being (IMR lag)	0.199	0.312	-0.660	0.286 *	0.622	0.365
Political rights (lag)	0.451	0.159 **	0.051	0.092	0.377	0.162 *
Capital city grid cell	0.964	0.455 *	1.539	0.353 **	0.747	0.493
Distance to border (ln)	-0.163	0.158	-0.185	0.113	-0.171	0.128
Non-violence media trend (ln)	0.399	0.133 **	0.169	0.081 *	0.409	0.111 **
Random part						
Country level σ^2	1.1285		1.0473		1.3798	
Model diagnostics						
AUC-ROC	0.8715		0.9233		0.8318	
AUC-PR	0.1425		0.1533		0.0894	
Brier score	0.0171		0.0078		0.0197	

Significance codes: ** $p < 0.01$, * $p < 0.05$. σ^2 is variance.

Number of observations (grid-months) = 816,552.

Random effect intercepts not reported for 42 countries.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

To further test the robustness of our model, we replace the ACLED dataset with the Uppsala Conflict Data Program Georeferenced Event Data (UCDP-GED) data (Croicu & Sundberg, 2015; Sundberg & Melander, 2013). Version 3.0 of these data covers the period 1989-2014. These data are based on a different coding methodology than ACLED, and so are not directly comparable. In general, they are more conservative than ACLED in capturing violence, requiring a minimum of 25 battle related deaths per year. UCDP-GED data do not include unidentified actors, for example, or events that take place outside of an “incompatibility” between actors that resulted in at least one death in a year. We excluded events with temporal precision that were only as precise as the year or longer, and for events that spanned

more than one day, we used the median day. We also removed events with poor spatial precision (i.e., geolocated to an entire country or large region of a country) from the analysis.

Table A2 shows a larger temperature and political rights effect for the UCDP-GED data. The capital city indicator is not significant, which reflects one of the differences from the ACLED data, because riots and protests tend to cluster in capitals and are not included in UCDP-GED. The UCDP-GED data are also known to be less sensitive to media reporting trends over time (Hegre, 2013), which is confirmed in the non-significance of this variable. To reflect this difference, we also estimate the model by excluding the non-violence media trends, which has the effect of increasing the temperature anomaly effect. Lastly, we replace the Freedom House political rights score with the “polity2” variable from the Polity IV Project (2014). Since it is coded differently from the political rights score, the negative sign is expected, but it weakens the overall model. We also substitute the polity2 term for the Freedom House measures in the analysis of ACLED data with similar results (Table A3).

Table A2. UCDP-GED models (1989-2012) with media trend and polity variants.

	UCDP-GED		UCDP, no media trend		UCDP, polity	
	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>
Fixed part						
Constant	-18.365	3.117 **	-16.390	3.430 **	-13.771	2.846 **
Precipitation (SPI6)	0.030	0.040	0.053	0.050	0.024	0.064
Temperature (TI6)	0.155	0.052 **	0.196	0.060 **	0.211	0.062 **
Population (ln)	0.785	0.200 **	0.807	0.198 **	0.794	0.206 **
Well-being (IMR lag)	0.352	0.392	0.051	0.307	0.182	0.319
Political rights (lag)	0.579	0.203 **	0.469	0.201 *		
Polity (lag)					-0.028	0.061
Capital city grid cell	0.417	0.689	0.390	0.635	0.409	0.704
Distance to border (ln)	-0.209	0.156	-0.207	0.124	-0.204	0.130
Non-violence media trend (ln)	0.271	0.168				
Random part						
Country level σ^2	2.8841		3.2283		3.5983	
Model diagnostics						
AUC-ROC	0.8362		0.8335		0.8309	
AUC-PR	0.0928		0.0935		0.0884	
Brier score	0.0185		0.0186		0.0188	

Significance codes: ** $p < 0.01$, * $p < 0.05$. σ^2 is variance.

Number of observations (grid-months) = 593,856.

Random effect intercepts not reported for 41 countries.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

There is evidence that drought is best measured using a combination of temperature and precipitation metrics (AghaKouchak et al., 2014). To test this, we calculate a combined hot and dry

anomaly metric by subtracting SPI6 from TI6 such that months that are both unusually hot and dry will be highlighted. This “hot & dry” model (Table A3) does not perform any better than our main models (Table 2). We further test temperature and precipitation combinations by calculating a 6-month Standardized Precipitation Evapotranspiration Index (SPEI6) using a 30-year rolling climatology. This index is calculated using the difference between precipitation and potential evapotranspiration (PET), where PET is modeled using the Thornthwaite equation which considers observed temperature and the latitude of each grid cell (Vicente-Serrano et al. 2010). Table A3 shows the results for models using the SPEI6 as the only climate metric and when combined with TI6. The estimates indicate that adding the combined evapotranspiration index does not improve the predicative capability of the model.

Table A3. ACLED model variants for polity score and drought metrics

	Polity		Hot & dry		SPEI6		TI6 & SPEI6	
	Estimate	SE	Estimate	SE	Estimate	SE	Estimate	SE
Fixed part								
Constant	-14.419	1.532 **	-16.640	2.186 **	-16.634	2.183 **	-16.624	1.967 **
Precipitation (SPI6)	0.029	0.030						
Temperature (TI6)	0.125	0.046 **					0.113	0.051 *
Hot & dry (TI6 - SPI6)			0.049	0.027				
Evapotranspiration (SPEI6)					-0.038	0.032	0.021	0.034
Population (ln)	0.822	0.138 **	0.817	0.143 **	0.817	0.139 **	0.819	0.129 **
Well-being (IMR lag)	0.377	0.361	0.383	0.284	0.388	0.292	0.381	0.228
Political rights (lag)			0.340	0.135 *	0.342	0.136 *	0.337	0.116 **
Polity (lag)	-0.012	0.039						
Capital city grid cell	1.003	0.371 **	1.000	0.394 *	0.996	0.384 **	1.004	0.395 *
Distance to border (ln)	-0.143	0.114	-0.145	0.117	-0.145	0.128	-0.147	0.115
Non-violence media trend (ln)	0.299	0.145 *	0.390	0.086 **	0.391	0.087 **	0.380	0.081 **
Random part								
Country level σ^2	0.8890		0.8902		0.8964		0.8819	
Model diagnostics								
AUC-ROC	0.8441		0.8452		0.8450		0.8454	
AUC-PR	0.1545		0.1582		0.1582		0.1582	
Brier score	0.0383		0.0379		0.0379		0.0379	

Significance codes: ** p < 0.01, * p < 0.05. σ^2 is variance.

Number of observations (grid-months) = 816,552.

Random effect intercepts not reported for 42 countries.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

Lastly, we include variants of the main Model 1 and Model 2 from Table 2 that use a different calculation for the temperature and precipitation anomalies. Instead of using a 30-year rolling climatology, these measures use a static, long-term climatology from 1949-2012. The estimates here

(Table A4) are nearly identical to those presented in Table 2, with the temperature effect slightly stronger using the long-term climatology as the reference period.

Table A4. Variants of the main Model 1 and Model 2 using the long-term climatology, 1949-2012, for calculating SPI6 and TI6.

	Constant Climatology		Constant Clim., no SPI6	
	Estimate	SE	Estimate	SE
Fixed part				
Constant	-16.578	2.167 **	-16.585	2.084 **
Precipitation (SPI6)	0.018	0.028		
Temperature (TI6)	0.112	0.046 *	0.109	0.050 *
Population (ln)	0.819	0.134 **	0.819	0.133 **
Well-being (IMR lag)	0.403	0.273	0.402	0.260
Political rights (lag)	0.342	0.128 **	0.342	0.125 **
Capital city grid cell	1.000	0.377 **	1.000	0.341 **
Distance to border (ln)	-0.144	0.120	-0.144	0.122
Non-violence media trend (ln)	0.368	0.094 **	0.369	0.089 **
Random part				
Country level σ^2	0.8967		0.8969	
Model diagnostics				
AUC-ROC	0.8453		0.8453	
AUC-PR	0.1577		0.1577	
Brier score	0.0380		0.0380	

Significance codes: ** $p < 0.01$, * $p < 0.05$. σ^2 is variance.

Number of observations (grid-months) = 816,552.

Random effect intercepts not reported for 42 countries.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

Country-level random effects

This section presents the country-level random effects of the main models. Table A5 shows the random effects for each country labeled by its 3-letter country code. Due to the hierarchical resampling technique, we calculated the standard errors for the random effects portion from a subset of the 200 bootstrapped models. The number of models varies for each country, but averages about 127. Figure A5, sometimes referred to as a caterpillar plot, shows the ranked random effects for Model 2 using the same bootstrapped standard errors reported in Table A5. Low ranks (e.g. South Sudan and Tanzania) indicate that the main fixed effects model overpredicts violence and is corrected downward by the

random effect coefficient. Similarly, large random effects (e.g. Somalia and Namibia) indicate the model underpredicts violence for those countries.

Table A5. Full coefficients for Models 1 & 2.

	Model 1		Model 2, no SPI6	
	Estimate	SE	Estimate	SE
Fixed part				
Constant	-16.628	2.089 **	-16.638	2.205 **
Precipitation (SPI6)	0.028	0.032		
Temperature (T16)	0.108	0.042 *	0.104	0.041 *
Population (ln)	0.820	0.130 **	0.819	0.140 **
Well-being (IMR lag)	0.382	0.267	0.381	0.282
Political rights (lag)	0.337	0.123 **	0.337	0.132 *
Capital city grid cell	1.004	0.354 **	1.003	0.379 **
Distance to border (ln)	-0.147	0.118	-0.146	0.113
Non-violence media trend (ln)	0.379	0.082 **	0.381	0.082 **
Random part				
AGO	0.636	0.392	0.636	0.389
BDI	1.416	0.273 **	1.416	0.263 **
BEN	-1.384	0.374 **	-1.382	0.424 **
BFA	-0.970	0.592	-0.968	0.405 *
BWA	0.615	0.556	0.609	0.565
CAF	0.782	0.542	0.781	0.578
CIV	-0.134	0.773	-0.130	0.505
CMR	-1.610	0.882	-1.612	0.344 **
COG	-0.821	0.309 **	-0.823	0.922
DJI	1.174	0.381 **	1.179	0.822
ETH	-0.479	0.355	-0.481	0.420
GAB	-0.188	1.412	-0.185	0.823
GHA	-0.824	0.246 **	-0.827	1.261
GIN	-0.996	0.366 **	-0.996	1.423
GNB	0.602	1.392	0.604	1.337
GNQ	-0.606	0.436	-0.599	1.553
KEN	0.306	0.312	0.303	0.374
LBR	1.176	0.391 **	1.179	0.389 **
LSO	0.540	0.516	0.538	0.521
MLI	-0.430	1.922	-0.430	0.503
MOZ	0.682	0.295 *	0.680	0.286 *
MRT	0.092	2.092	0.099	0.619
MWI	-1.117	0.321 **	-1.121	0.411 **
NAM	1.710	0.480 **	1.709	0.527 **
NER	-0.762	2.313	-0.758	0.338 *
NGA	-1.220	0.345 **	-1.220	0.302 **
RWA	-0.018	0.538	-0.015	0.496
SDN	0.229	0.339	0.228	2.587
SEN	0.058	0.556	0.063	0.552
SLE	0.912	0.500	0.915	2.595
SOM	1.540	2.641	1.540	2.634
SWZ	0.347	0.315	0.344	0.384
TCD	-0.127	0.331	-0.121	0.321
TGO	-0.635	0.387	-0.629	0.410
TZA	-1.639	3.161	-1.643	0.287 **
UGA	0.536	0.348	0.540	3.423
ZAF	1.484	0.368 **	1.474	0.338 **
ZAR	-0.428	3.301	-0.430	3.328
ZMB	0.022	0.350	0.022	0.371
ZWE	1.460	3.407	1.457	0.462 **
ERI	-0.053	3.793	-0.058	3.660
SSD	-1.794	0.577 **	-1.803	0.604 **
Model diagnostics				
Country level σ^2	0.8824		0.8828	
AUC-ROC	0.8454		0.8454	
AUC-PR	0.1582		0.1581	
Brier score	0.0379		0.0379	

Significance codes: ** p < 0.01, * p < 0.05. σ^2 is variance.

Number of observations (grid-months) = 816,552.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

SE = bootstrap standard errors with 200 hierarchical resamples.

Random effects SEs calculated from ~127 coefficients due to hierarchical resampling.

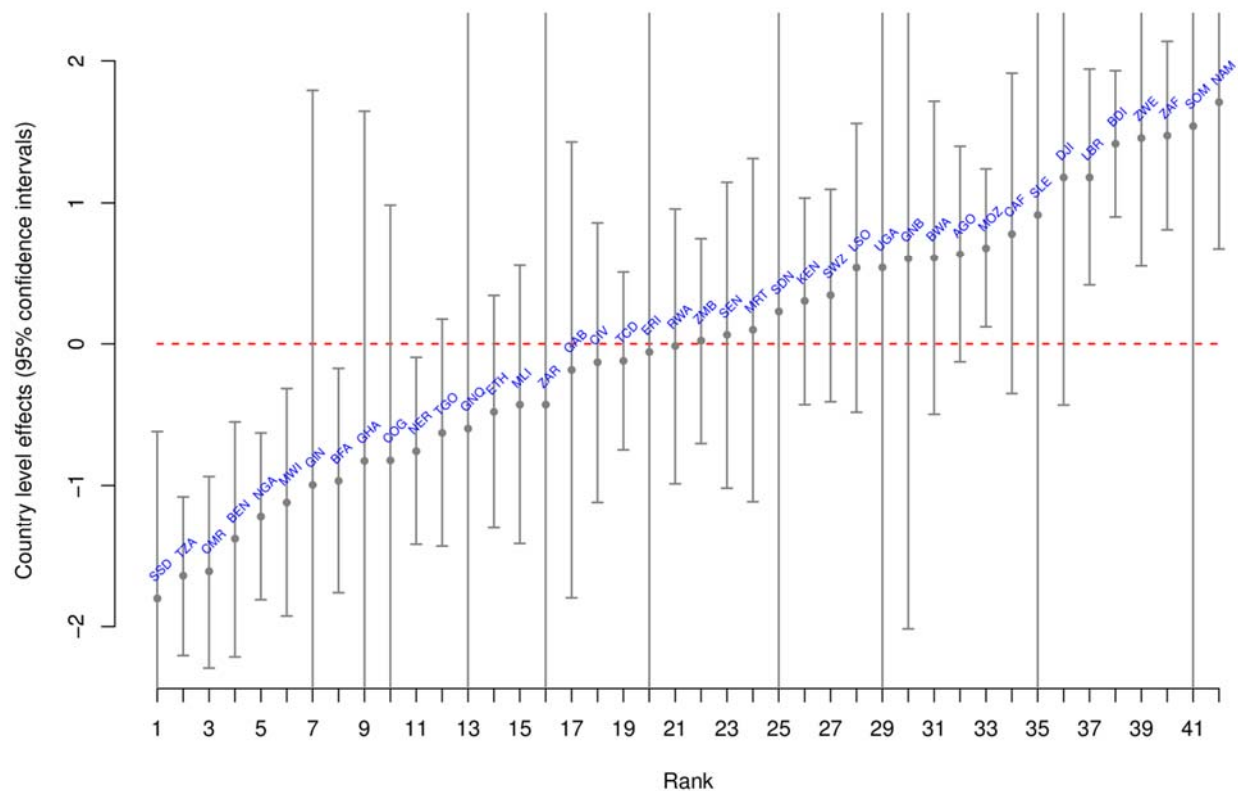


Figure A5. Country-level effects for Model 2. Grey bars show 95% confidence interval using bootstrapped standard errors.

Model validation

To validate our model’s ability to forecast violence, we test our predictive capability on a subset of the data and compare in-sample and out-of-sample model diagnostics. Table A6 shows the estimates for Model 2 for the temporal subset, 1980-2007. We then use these coefficients to predict violence in the subsequent five years, 2008-2012, and compare those predicted values to the observed values. As expected, the predictive capability declines for the out-of-sample data (e.g. AUC-ROC drops from 0.8487 in-sample to 0.7881), but remains high enough to maintain confidence in the model’s ability to forecast future violence.

Table A6. Out-of-sample validation for Model 2.

Fixed part	Model 2, 1980-2007	
	Estimate	SE
Constant	-17.242	2.571 **
Temperature (TI6)	0.102	0.051 *
Population (ln)	0.882	0.166 **
Well-being (IMR lag)	0.524	0.234 *
Political rights (lag)	0.321	0.135 *
Capital city grid cell	0.829	0.365 *
Distance to border (ln)	-0.119	0.138
Non-violence media trend (ln)	0.357	0.076 **
Random part		
Country level σ^2	0.9774	
Model diagnostics	<i>In-sample,</i> 1980-2007	<i>Out-of-sample,</i> 2008-2012
AUC-ROC	0.8487	0.7881
AUC-PR	0.1434	0.1676
Brier score	0.0331	0.0640

Significance codes: ** $p < 0.01$, * $p < 0.05$. σ^2 is variance.

Number of observations (grid-months) = 692,832.

Random effect intercepts not reported for 41 countries.

SE = bootstrap standard errors with 200 hierarchical resamples.

AUC = area under the curve; ROC = receiver operator characteristic.

PR = precision-recall.

We further validate the model by replacing the observed CRU climate data with the six historical climate simulations forced using observed SSTs (Figure 2). These additional models (Table A7) are limited to the period June 1980 through 2010 due to climate simulation data constraints. To assist in the comparison, Model 2 is reported in Table A7 constrained to the matching time period. The temporal constraints of the climate simulations also mean that the precipitation and temperature anomalies were calculated using the entire set of observations, not a rolling 30-year period. Though the overall simulated temperature trends match the observed trend (Figure 2), the spatial distribution of anomalies differs enough from the observed record to reduce the temperature effect, leaving it statistically significant in only one of the specified SST simulations (SST6 in Table A7). This result means that temperature effects on violence is not only a function of continent-wide warming – *it matters where and when it gets warmer*. This result also suggests that the specific monthly spatial variation of forecast violence that is attributable to hotter temperatures is not reliable, but that the overall temperature-related violence estimate is plausible.

Table A7. Model validation using six specified sea surface temperature (SST) simulations for the historical climate model inputs.

	Model 2, June 1980-2010		SST1		SST2		SST3		SST4		SST5		SST6	
Fixed part	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>	<i>Estimate</i>	<i>SE</i>
Constant	-16.584	2.331 **	-16.568	2.465 **	-16.648	2.443 **	-16.633	2.462 **	-16.657	2.456 **	-16.669	2.513 **	-16.591	2.269 **
Temperature (TI6)	0.127	0.060 *	0.070	0.062	-0.012	0.046	0.005	0.044	-0.016	0.048	-0.028	0.054	0.078	0.037 *
Population (ln)	0.831	0.148 **	0.827	0.164 **	0.827	0.156 **	0.827	0.159 **	0.827	0.156 **	0.827	0.159 **	0.827	0.142 **
Well-being (IMR lag)	0.501	0.256	0.484	0.277	0.483	0.279	0.483	0.302	0.483	0.285	0.482	0.305	0.486	0.271
Political rights (lag)	0.344	0.131 **	0.349	0.141 *	0.347	0.133 **	0.348	0.138 *	0.348	0.142 *	0.348	0.138 *	0.349	0.131 **
Capital city grid cell	0.949	0.369 *	0.945	0.426 *	0.946	0.387 *	0.946	0.411 *	0.946	0.388 *	0.946	0.384 *	0.945	0.353 **
Distance to border (ln)	-0.141	0.122	-0.141	0.135	-0.142	0.133	-0.142	0.133	-0.142	0.132	-0.142	0.132	-0.140	0.128
Non-violence media trend (ln)	0.342	0.091 **	0.355	0.102 **	0.374	0.088 **	0.370	0.097 **	0.375	0.089 **	0.378	0.092 **	0.358	0.082 **
Random part														
Country level σ^2	0.9087		0.9007		0.8976		0.8986		0.8974		0.8976		0.9044	
Model diagnostics														
In-sample AUC-ROC	0.8454		0.8451		0.8448		0.8448		0.8447		0.8446		0.8449	
Out-of-sample AUC-ROC, 2011-12	0.8018		0.7990		0.7994		0.7993		0.7994		0.7995		0.7992	
In-sample AUC-PR	0.1495		0.1497		0.1495		0.1496		0.1497		0.1495		0.1491	
Out-of-sample AUC-PR, 2011-12	0.2175		0.2118		0.2113		0.2116		0.2113		0.2088		0.2121	
In-sample Brier Score	0.0352		0.0352		0.0352		0.0352		0.0352		0.0353		0.0352	
Out-of-sample Brier, 2011-12	0.0675		0.0680		0.0693		0.0690		0.0694		0.0696		0.0681	

Significance codes: ** p < 0.01, * p < 0.05. σ^2 is variance.

Number of observations (grid-months) = 756,754.

Random effect intercepts not reported for 41 countries.

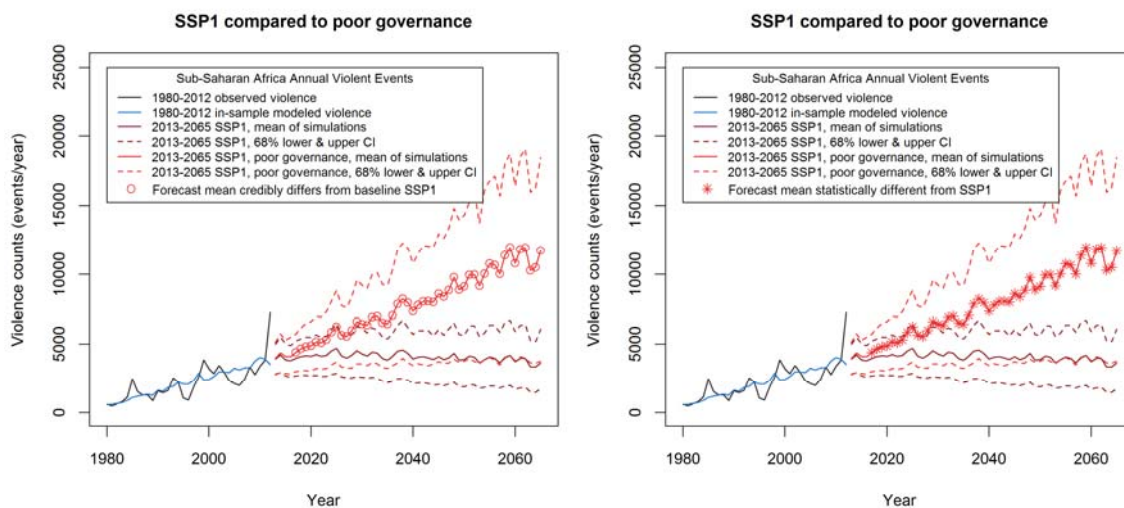
SE = bootstrap standard errors with 200 hierarchical resamples.

AUC = area under the curve; ROC = receiver operator characteristic; PR = precision-recall.

Conflict forecasts differences evaluation

We expand on the evaluation of conflict forecasts differences by considering whether pairs of forecasts differ for a given year. Though the 68% confidence intervals for the SSP1 forecasts and their modified versions overlap (Figure A6), this visual method is not reliable for judging meaningful differences (Schenker & Gettleman, 2001). Instead, we formally test for differences using both a Bayesian difference of means analysis (Kruschke 2013) and a null hypothesis significance test, the Welch two sample t test. For these tests, we use a 99% highest density interval (HDI) and corresponding p -value threshold of 0.01. For the Bayesian analysis, a Markov chain Monte Carlo (MCMC) sampling procedure is used to generate the posterior distribution of parameter values. For the two input forecast groups, their means, standard deviations, and a shared normality parameter are used as inputs to generate 100,000 MCMC samples for these five parameters. The analysis uses noncommittal priors to prevent introducing any bias to the posterior distribution.

Figure A6 shows general agreement between these two tests, with the t test detecting more differences in the forecasts than the Bayesian method. They both show that the effects of poor governance, when compared to the SSP1 scenario of improving governance, would result in substantially more violence within a matter of years. The effects of high fertility when compared against the low fertility of SSP1 are not noticeably different until about 2040. In contrast to the effects of governance and population growth, the effects of higher temperatures (RCP 8.5) compared to the lower warming scenario of RCP 2.6 show only sporadic meaningful differences, with no consistent effect.



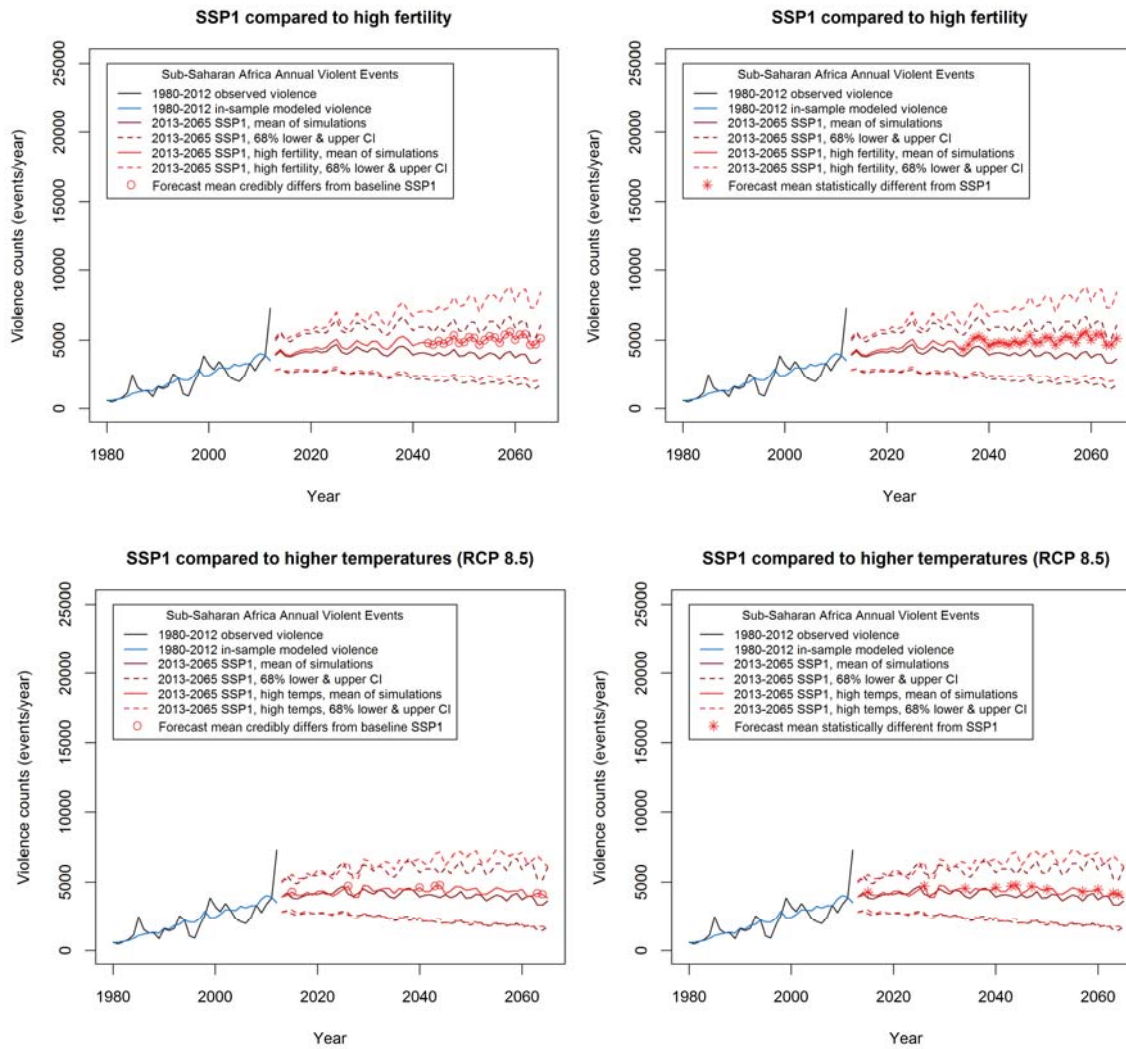


Figure A6. Plots show 68% confidence intervals (CI) for comparing SSP1 to forecasts with poor governance, high fertility, and higher temperatures. The circular points in the left column highlight years whose 200 simulated SSP1 lower and upper forecasts differ credibly using a 99% highest density interval (HDI). Stars in the right column highlights years whose t tests are statistically significant for a p -value < 0.01 .

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